Statistical Analysis with R (Estimation of binary logit and probit models)

We use the data on the existence of capital punishment system in U.S. states in "dpdata.txt". When

you open dpdata.txt you see:

# Data on the existence of capital punishment system in U.S. states										
# D1: Existence of capital punishment, D1=1: existence, D2=0: non-existence										
# T : median length of stay in prison for murderers released from prison										
# Y : median of household income										
# NW: proportion of non-Caucasian people										
D1	Т	Y	NW							
1	47	1.1	0.321							
1	58	0.92	0.224							
1	82	1.72	0.127							
1	100	2.18	0.063							
1	222	1.75	0.021							
1	164	2.26	0.027							
1	161	2.07	0.139							
1	70	1.43	0.218							
1	219	1.92	0.008							
0	81	1.82	0.012							
1	209	2.34	0.076							

Our aim is to analyze what variable affects the existence of capital punishment in U.S. After changing the working directory to the one in which the dpdata.txt is saved, you can load the data by typing:

> dpdata=read.table("dpdata.txt",header=T,skip=5)

To see the first 8 rows, type

.....

> head(dpdata,8)

then you'll see

D1		Т	Y	NW
1	1	47	1.10	0.321
2	1	58	0.92	0.224
3	1	82	1.72	0.127
4	1	100	2.18	0.063
5	1	222	1.75	0.021
6	1	164	2.26	0.027
7	1	161	2.07	0.139
8	1	70	1.43	0.218

2. Esimating Logit Model

We let D1 be dependent variable and other variables be explanatory variables. We estimate the parameters β_0 , β_1 , β_2 , and β_3 of

 $Pr(D1_i=1) = F(\beta_0 + \beta_1 T_i + \beta_2 Y_i + \beta_3 NW_i),$

where F() is the distribution function of logistic distribution. For this purpose, we use R-function glm(). Type

```
> resultlogit=glm(D1~T+Y+NW, family=binomial(link="logit"), data=dpdata)
```

Then, you'll see (it is no problem to ignore the warning message)

```
Warning message:
glm.fit: fitted probabilities numerically 0 or 1 occurred
>
```

Note that in the case that you use all variables other than D1 in dpdata as explanatory variables, you can omit explanatory variables by typing "."

> resultlogit=glm(D1~.,family=binomial(link="logit"),data=dpdata)

Estimation result can be checked by

```
> summary(resultlogit)
Call:
glm(formula = D1 ~ ., family = binomial(link = "logit"), data = dpdata)
Deviance Residuals:
   Min 1Q Median
                              3Q
                                     Max
-2.52498 0.00000 0.03012 0.34398 1.65742
Coefficients:
         Estimate Std. Error z value Pr(>|z|)
(Intercept) -11.875982 7.427891 -1.599 0.1099
           0.016051 0.009989 1.607 0.1081
Т
           3.941698 3.296000
                               1.196
Y
                                       0.2317
          91.021874 36.347643 2.504 0.0123 *
NW
Signif. codes: 0 `***' 0.001 `**' 0.01 `*' 0.05 `.' 0.1 ` ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 44.584 on 43 degrees of freedom
Residual deviance: 21.166 on 40 degrees of freedom
AIC: 29.166
```

Number of Fisher Scoring iterations: 9

The values in column of Estimate are the estimates of coefficients, those in the column of Std. Error are standard errors of those estimates, and z value is corresponding to t value in regression analyses for examining the null hypothesis that the true value of each coefficient is 0, Pr(> |z|) is P-value of z. AIC is the value called "Akaike Information Criteria" which is a measure of fit of model. Smaller AIC implies better model.

Using these coefficient estimates, we predict the probability of $D1_i$ being 1 (it is like a fitted value in regression analyses) by

> resultlogit\$fitted

You'll see (the values whose actual results are 0 are highlighted in yellow)

1	2	3	4	5	6
1.00000000	0.99999789	0.99958202	0.98298256	0.62167664	0.89303614
7	8	9	10	11	12
0.99999007	0.99999960	0.48383649	<mark>0.09031995</mark>	0.99950961	1.0000000
13	14	15	16	17	18
0.86630986	0.80140149	1.00000000	0.71349310	0.99999724	<mark>0.17351756</mark>
19	20	21	22	23	24
<mark>0.12356981</mark>	0.98906380	1.00000000	<mark>0.41611966</mark>	0.99999999	0.52272045
25	26	27	28	29	30
<mark>0.95873632</mark>	0.99745346	0.94399363	0.99496946	0.99780858	0.99390984
31	32	33	34	35	36
0.98256814	0.25321599	0.98127178	<mark>0.42999810</mark>	1.00000000	0.93813672
37	38	39	40	41	42
0.99986822	0.99823730	<mark>0.35017066</mark>	0.99999989	0.86341648	<mark>0.21768336</mark>
43	44				
0.91007074	0.51136739				

You can compare these probabilities with actual observations.

```
> dpdata$D1
[1] 1 1 1 1 1 1 1
[7] 1 1 1 1 0 1 1
[13] 1 1 1 1 0
[19] 0 1 1 0 1 1
[25] 0 1 1 1 0 1 1
[31] 1 1 1 0 1 1
[37] 1 1 0 1 1 0
[43] 1 0
```

You can also check what other results you can see by typing

>help(glm)

3. Calculating Marginal Probability Effect

Here, we calculate MPE using R. glm() function does not automatically calculate MPE, so we have to calculate it by ourselves using the formula. In the case of the logit model, it is relatively easily calculated. Let $\Lambda(x)$ be the distribution function of logistic distribution. Remember that, then, the MPE is

$$MPE_{ij} = f_{\text{logit}}(\beta_0 + \beta_1 T + \beta_2 Y + \beta_3 NW) \beta_j$$

```
= \Lambda(\beta_0 + \beta_1 T + \beta_2 Y + \beta_3 NW) [1 - \Lambda(\beta_0 + \beta_1 T + \beta_2 Y + \beta_3 NW)] \beta_j.
```

where $f_{\text{logit}}(.)$ is the pdf of the logit distribution. Thus, for example, for i = 4 and j = 3, MPE_{ij} can be calculated by:

> Lam=resultlogit\$fitted

(Remember this is the value of $\Lambda(\beta_0 + \beta_1 T + \beta_2 Y + \beta_3 NW)$ in the logit model)

> coeflogit=coef(resultlogit)

(coef(result1) returns a vector of coefficient estimates)

> MPElogit=Lam*(1-Lam)*coeflogit[3]

("*" is element-by-element product of matrix(or vector))

> MPElogit43=MPElogit[4]

(MPE[k] is k-th element of MPE)

The answer is

```
> MPElogit43
4
0.06593612
```

The AMPE is the average of MPE, so it can be calculated by

```
> AMPElogit = mean(MPElogit)
> AMPElogit
[1] 0.3045834
```

4. Estimating Probit Model

Similarly, we can estimate probit model too.

```
>resultprobit=glm(D1~.,family=binomial(link="probit"),data=dpdata)
>summary(resultprobit)
Call:
glm(formula = D1 ~ ., family = binomial(link = "probit"), data = dpdata)
Deviance Residuals:
    Min
         1Q Median
                               3Q
                                     Max
-2.33906 0.00000 0.00362 0.34356 1.57845
Coefficients:
          Estimate Std. Error z value Pr(>|z|)
(Intercept) -6.750043 4.041732 -1.670 0.09490 .
          0.007607 0.005215 1.459 0.14468
т
          2.342983 1.820228 1.287 0.19803
Y
          51.802479 19.215316 2.696 0.00702 **
NW
Signif. codes: 0 `***' 0.001 `**' 0.01 `*' 0.05 `.' 0.1 ` ' 1
(Dispersion parameter for binomial family taken to be 1)
```

Null deviance: 44.584 on 43 degrees of freedom Residual deviance: 21.402 on 40 degrees of freedom AIC: 29.402 Number of Fisher Scoring iterations: 10 The only difference is that link="logit" changes to link="probit". Again only NW is significant at 5% nominal level.

5. Alternative Way to Compute MPE

In Section 3, we used a well known relationship between the distribution function and the probability density function of logistic distribution. In R, there is a function dlogis(x), which returns the value of pdf of the logistic distribution at x (if x is a vector, then dlogis(x) returns a vector of the values of pdf at x). In this section, we calculate the MPE with a logit model, using this function. This way of calculation can be easily extended to the probit model.

Recall that MPE of *j*-th explanatory variable for *i*-th individual is given by

 $MPE_{ij} = f_{\text{logit}}(\beta_0 + \beta_1 T_i + \beta_2 Y_i + \beta_3 NW_i) \beta_j.$

 β_j was already obtained as coef(j). What we have to calculate is the value of $\beta_0 + \beta_1 T_i + \beta_2 Y_i + \beta_3 NW_i$. We calculate this value in the following way (this is a little bit complicated and needs some additional knowledge about R).

Consider a matrix of explanatory variables such as

$$\mathbf{X} = \begin{bmatrix} 1 & T_1 & Y_1 & NW_1 \\ 1 & T_2 & Y_2 & NW_2 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & T_n & Y_n & NW_n \end{bmatrix},$$

where, for example, T_i is a value of T for *i*-th individual, etc. Then the value of $\beta_0 + \beta_1 T_i + \beta_2 Y_i + \beta_3 NW_i$ is *i*-th row of X β , where $\beta = [\beta_0, \beta_1, \beta_2, \beta_3]$, where A' is the transpose of a matrix(or vector) A. Below, we calculate X β .

First, noting that the number of observation (n) is 44, we construct the first column of X (that is n by 1 vector of ones)

> const=numeric(44)+1

Here numeric (n) returns a vector of zeros of length of n. "a +1" adds one to all components of a vector a. Then, matrix X can be constructed by typing

> X=cbind(const,dpdata[,2:4])

Here cbind(A, B) combines two matrices A and B so that cbind(A, B) = [A, B], and A[, a:b] extracts a- to b-th columns of a data object A.

Next, XB is calculated by typing

> XB=as.matrix(X)%*%as.matrix(coeflogit,4,1)

Here, as.matrix(A) changes a data object A to a matrix object, and as.matrix(b, c, d) changes a vector object b to a c by d matrix object so that we can apply matrix calculation to these objects(here R recognizes as.matrix(X) as 44 by 4 matrix and as.matrix(coeflogit, 4, 1) as 4 by 1 matrix(vector) so that we can apply the operator "%*%", which is for usual matrix product).

Finally, using dlogis(.) function, we have

> flogis=dlogis(XB)

Note that this is equal to Lam* (1-Lam) that we calculated in Section 3. Thus, we can calculate MPE_{43} , by

- > MPElogit2=flogis*coeflogit[3]
- > MPElogit2_43=MPElogit2[4]
- > MPElogit2_43
- [1] 0.06593612

We can see that we obtain the same value in Section 3.

Similarly, we can calculate the average MPE by

```
> AMPElogit=mean(MPElogit2)
```

- > AMPElogit
- [1] 0.3045834

Exercise

R-function "dnorm(x, mean=m, sd=s)" returns the value of the pdf of normal distribution with mean m and standard deviation s at x. Using this function, calculate the MPE for i = 4 and j = 3, for the probit model. Similarly, calculate the average MPE.