

To appear in *Phonology*
This is a near-final version.

Token frequency modulates optional paradigm uniformity in Japanese voiced velar nasalization

June 26, 2024

Abstract

This paper explores the role of token frequency and global prosodic length in conditioning optional paradigm uniformity in Japanese voiced velar nasalization, with data from two *wug*-tests carried out with speakers of Tōhoku Japanese. Experiment 1 demonstrates that frequency-conditioning observed in corpus data is reproduced in existing and novel compounds, and holds at the level of the speaker. Experiment 2 focuses on a typologically-unusual pattern where overall compound length in mora seems to influence nasalization, a candidate for a “counting” pattern in phonology. We find that instead of overall length, speakers are sensitive to the length of the second member of the compound, undercutting the viability of the mora-counting analysis. We discuss the importance of the results in adjudicating between existing models of how token frequency impacts the phonological grammar, and suggest that only theories that allow individual morphemes to exhibit frequency-sensitive behavior are sufficiently expressive to model the finding.

1 Introduction

This paper explores the role of token frequency in conditioning optional paradigm uniformity in Japanese voiced velar nasalization (henceforth, simply “nasalization”). We begin with a review of the salient findings of Breiss et al. (2021)’s corpus study of Yamanote Japanese, where the probability of nasalization is influenced by overall compound frequency, the frequency of the second element (N2), and, unexpectedly, the total number of mora in the compound. We then report two *wug*-tests with speakers of the phonologically-conservative Tōhoku dialect of Japanese (spoken in northern parts of the main island of Japan). The first experiment demonstrates that the frequency-conditioning observed in corpus data is reproduced in existing and novel compounds, and holds at the level of the individual speaker. The second experiment focuses specifically on whether, after controlling for frequency, overall compound length influences nasalization, and

25 finds that it does not, instead revealing sensitivity to only the size of N2, regardless of the size
26 of N1 (the first element in the compound). We conclude by discussing the implications of these
27 findings for existing theories of token frequency in the grammar.

28 **1.1 Token frequency and phonological patterns**

29 The influence of lexical frequency on phonological patterning has been much debated in the
30 literature. In the classic generative tradition, frequencies—or statistical information in general—
31 are considered to lie outside of grammatical competence, as argued for example in Chomsky’s
32 *Syntactic Structures* (Chomsky, 1957). The classic argument given in that book was that the tran-
33 sitional probability in English from *fragile* to *whale* is plausibly zero, but this sequence does not
34 make a sentence containing it ungrammatical. The *Sound Pattern of English* (Chomsky and Halle,
35 1968) (SPE) and much subsequent work that followed its tradition did not seriously consider the
36 effects of lexical frequency or other statistical information on phonological patterns either (Co-
37 etzee and Kawahara, 2013). Phonological rules, as formulated in the SPE style, were not sensitive
38 to usage frequencies. More generally, effects related to statistical information, including usage
39 frequency were, implicitly or explicitly, considered to be a matter of performance and not com-
40 petence (Newmeyer, 2003) (see the reviews presented in Coetzee and Kawahara 2013 and Shaw
41 and Kawahara 2018).

42 On the other hand, there was a recurrent observation that for example, deletion of a phono-
43 logical segment is more likely in frequent words than in non-frequent words (Bybee, 1999). This
44 observation was made as early as Hooper (1976), who observed that schwa is more likely to be
45 deleted in frequent words like *memory* than in non-frequent but phonologically-similar words like
46 *mammary*. A perhaps more well-studied phenomenon is [t/d]-deletion in English, which is very
47 common in frequent words like *and* and *list*, but not as very common in less frequent words like
48 *mast* or *jest* (Coetzee and Kawahara, 2013). Likewise, the experiment by Kawahara (2011) demon-
49 strates that OCP-induced geminate devoicing in Japanese loanwords is judged to be more natural
50 in words with higher lexical frequencies. File-Muriel (2009) argues that in some dialect of Span-
51 ish, weakening of /s/ occurs more likely in high frequency words than in low frequency words.
52 Zuraw (2007) examines frequency-conditioned application of markedness-reducing phonological
53 processes in a corpus of written Tagalog, and finds higher rates of repair within higher-frequency
54 units (words, clitic groups, etc.), subject to the markedness principles of the language.

55 Lexical frequency has also been found to be related to the degree a lexical item deviates from
56 grammar-wide trends in phonological alternations. Smith and Moore-Cantwell (2017) found that
57 higher-frequency comparative constructions are more likely to flout grammar-wide trends driven
58 by markedness. For example, the adjective *likely* itself forms the analytic comparative *more likely*
59 more than 99% of the time, while other phonologically-comparable, lower-frequency forms take

60 the analytic comparative only around 45% of the time. In a similar vein, Anttila (2006) and Mayer
61 (2021) found that higher-frequency morphologically-complex forms were more likely to behave
62 opaquely with respect to grammar-wide phonological processes.

63 These cases show that lexical frequency interacts with phonological patterns to a non-negligible
64 degree, and any adequate theory of phonology must have a means to model its effects. In usage-
65 based phonology (Bybee, 1999) as well as exemplar-theoretic phonology (Gahl and Yu, 2006), us-
66 age frequency is directly encoded in grammatical model. In the generative tradition too, Coetzee
67 and Kawahara (2013) argue that it is necessary—and possible—to incorporate the effects of lexical
68 frequency in the formal phonological grammar. Whether, and to what extent, lexical frequen-
69 cies affect phonological patterns, and which aspects of phonological patterns are (un)affected,
70 all remain important questions in current phonological research. In this paper, we provide ex-
71 perimental data bearing on this question, and discuss how different contemporary generative
72 models of lexicon-phonology interaction might model the role lexical token frequency has on
73 conditioning nasalization in both existing and novel compounds.

74 **2 Voiced Velar Nasalization**

75 **2.1 The basic patterns**

76 In many phonologically conservative dialects of Japanese, [ŋ] and [g] are allophonically dis-
77 tributed; here we summarize the generalizations in the literature about the Yamanote dialect,
78 a classic and conservative speaking style of the dialect spoken in the center of Tokyo (see Hibiya
79 (1995) for more on the sociolinguistic significance of nasalization). In dialects that exhibit nasal-
80 ization, /g/ is realized as [ŋ] in prosodic-word-medial position; e.g. /kagami/→[kaŋami] “mirror”
81 vs. /gimu/→[gimu] “obligation”.

82 This complementary distribution has been discussed extensively in the generative and pre-
83 generative literature on Japanese phonology (e.g. Kindaichi 1942; Trubetsky 1969; Labrune 2012).
84 Although properly a static phonotactic restriction, the prominence of compounding in Japanese
85 word-formation means that there are many contexts where the same morpheme can both surface
86 free-standing with initial [g], and as a second member of a compound (=N2) with initial [ŋ].
87 Thus there is ample opportunity to study status of the phonotactic restriction in the synchronic
88 grammar via the alternation it induces, as in (1)-(3). It is in this context that Ito & Mester treat
89 the phenomenon, first in Ito and Mester (1996) and later in Ito and Mester (2003), where they
90 formalize a constraint-based analysis of the alternation observed in compounds. Most relevant
91 for the current paper, they highlight the optionality of the alternation in cases where the second
92 member of the compound is also a free-standing word, as illustrated by the examples in (1)-(3).

- 93 (1) a. /hai + gan/ → [hai-ŋan] ~ [hai-gan]
 lung cancer
 94 “lung cancer”
- 95 b. /gan/ → [gan]
 96 cancer
 97 “cancer”
- 98 (2) a. /noo + geka/ → [noo-ŋeka] ~ [noo-geka]
 brain surgery
 99 “brain surgery”
- 100 b. /geka/ → [geka]
 101 surgery
 102 “surgery”
- 103 (3) a. /doku + ga/ → [doku-ŋa] ~ [doku-ga]
 poison moth
 104 “poison moth”
- 105 b. /ga/ → [ga], “moth”

106 The gist of their analysis is that the optionality is the result of two competing forces acting on
 107 the realization of the /g/-initial word that occurs in a compound as N2: (1) a paradigm uniformity
 108 effect to its base form (Steriade, 2000), which prefers [g] to [ŋ], and (2) a markedness constraint
 109 that favors nasalization in intervocalic positions, favoring [ŋ] to [g]. This analysis captures both
 110 the variability of compounds with free N2s, and also the obligatoriness of nasalization when N2
 111 is a bound morpheme, as in cases like (4).

- 112 (4) a. /doku + ga/ → [doku-ŋa], *[doku-ga]
 poison fang
 113 “poison fang”
- 114 b. /ga + ʒoo/ → [ga-ʒoo]
 fang castle
 115 “main castle”
- 116 c. /ga/ → *[ga] (a bound morpheme)
 fang
 117 “fang”

118 2.2 The corpus study of Breiss et al. (2021)

119 Breiss et al. (2021) carried out a quantitative investigation of the variability and optionality of
 120 nasalization, noted by Ito and Mester (1996). We reproduce Breiss et al.’s quantitative analysis

121 here, but refer the interested reader to the full text a more detailed treatment.

122 The data for the analysis was drawn from the 2016 NHK Pronunciation and Accent dictio-
123 nary (NHK, 1993) which represents a consensus view of expert dialectologists about normative
124 pronunciation in the Yamanote dialect, and was annotated with frequency information from the
125 Balanced Corpus of Contemporary Written Japanese (BCCWJ) (Maekawa et al., 2014). Figure 1
126 plots the number of words whose pronunciation was labeled as “undergoing” or “preferring to
127 undergo” nasalization in the dictionary, divided into whether N2s were bound morphemes (left
128 panel) or free-standing morphemes (right panel).¹

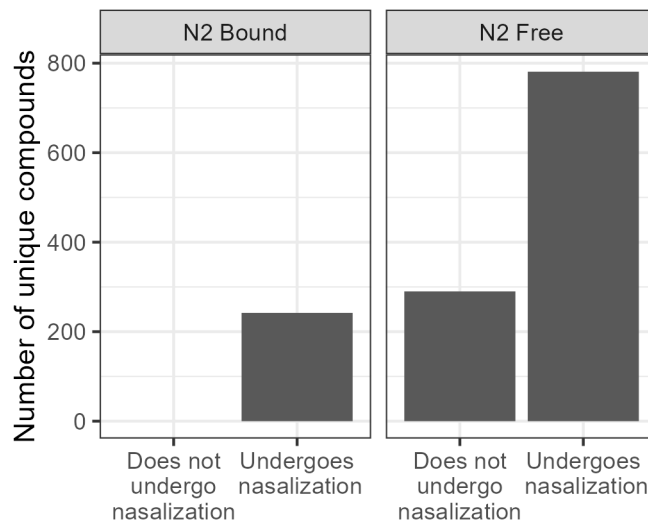


Figure 1: Division of the corpus of compounds according to whether a given compound undergoes (or prefers to undergo) nasalization or not (horizontal axis), divided by whether or not N2 is able to occur as a free form (panels). The vertical axis plots the number of unique compounds in each category.

129 The data in the left panel show that bound morphemes invariably undergo nasalization. The
130 data in the right panel support the claim of optionality in cases of a free-standing N2, offering
131 quantitative evidence supporting Ito and Mester (1996).

132 Turning next to the determinants of variation in compounds with free N2s, Breiss et al. (2021)
133 found the frequency of the whole compound and its second member to both be reliable predictors
134 of whether a given compound would undergo nasalization in dictionary data. These data are
135 plotted in Figure 2. The left panel shows that more frequent compounds are more likely to show
136 nasalization; the right panel shows, on the other hand, that compounds with more frequent N2s
137 are less likely to show nasalization.

¹Breiss et al. (2021) categorized morphemes appearing at least once as independent words in the BCCWJ corpus as free, and those not appearing independently as bound, with the understanding that this classification may not necessarily perfectly align with native speakers' intuition.

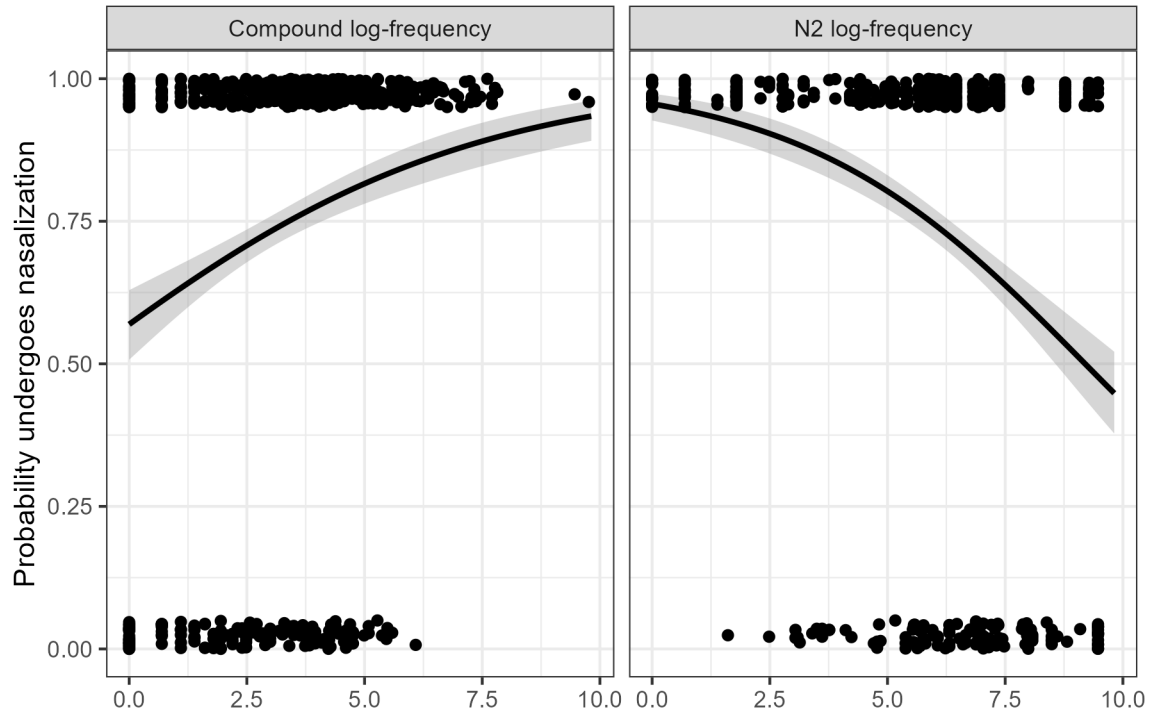


Figure 2: The effects of whole compound frequency (left panel) and N2 frequency (right panel) on the probability of nasalization, with binomial smooths. One dot represents one lexical item; vertical jitter has been added for readability.

138 Breiss et al. also found that the prosodic length of a compound was strongly related to whether
 139 it underwent nasalization, with shorter compounds being more likely to undergo, and longer
 140 compounds being more resistant to alternation. Figure 3 plots these findings.

141 Taken at face value, these data suggest that the synchronic grammar of the Yamanote dialect of
 142 Japanese exhibits a dependency between global prosodic length and a local segmental alternation
 143 which flies in the face of a often-cited claim that “phonology doesn’t count” (McCarthy and Prince
 144 1986 among many others). Though this is not as ironclad a generalization as it is often assumed
 145 to be (see in particular the arguments and data presented in Paster 2019), the trend in Figure
 146 3 is striking, with few typological parallels noted in the literature.² An alternative explanation
 147 not explored by Breiss et al., however, is that this relationship is illusory, a function of Zipf’s
 148 Law of Abbreviation (Zipf, 1935). The Law states that “the magnitude of words stands in an
 149 inverse (not necessarily proportionate) relationship to the number of occurrences” (Zipf, 1935,
 150 p.23); that is, more frequent words tend to be shorter. If we assume that this relationship also
 151 governs compounds in the Japanese lexicon, we would expect to find the pattern shown in Figure

²Paster (2019) notes that all the cases that involve counting are supra-segmental patterns, and thus, the pattern in Figure 3, if true, is indeed a surprising finding.

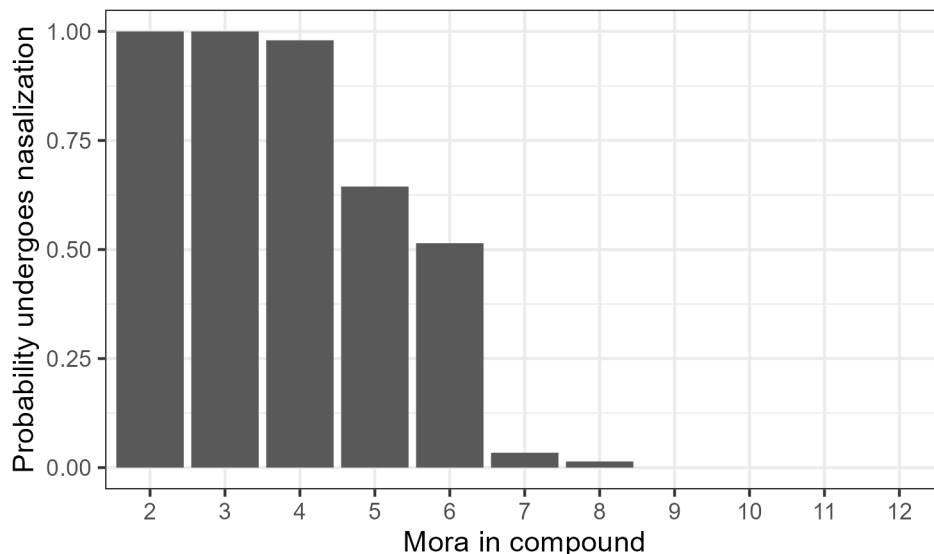


Figure 3: The effects of compound length in mora on the probability of nasalization in the corpus study of Breiss et al. (2021), reproduced with slightly different axis labels for consistency.

152 3, where compounds appear both shorter and more likely to undergo nasalization, both stemming
 153 from their high frequency.

154 As with all analysis of existing lexical items relying on corpus data, it is not clear how much
 155 evidence these data can provide about the synchronic grammars of speakers of such conservative
 156 dialects, and whether the frequency effects (functionally motivated) or the apparent length-effect
 157 (typologically very unusual) are in fact cognitively represented as such. Resolving these questions
 158 is of critical importance to how we construct our phonological theories, including whether (and
 159 if so, how) usage frequency of the morphemes impacts the the synchronic grammar.

160 3 Experiment 1

161 3.1 Aims

162 Our first experiment has three empirical goals. The first is to determine whether the optionality
 163 of the paradigm uniformity in existing compounds found in the corpus is operative at the level
 164 of the individual speaker. This is an important methodological point that is hard to resolve in
 165 corpus-based studies of variation, including Breiss et al. (2021), because it is possible that appar-
 166 ent variation in a corpus actually results from collapsing across different speakers with different
 167 categorical grammars.³ The second goal is to see whether the frequency-conditioning of this

³This is not to say that addressing the role of frequency within each individual speaker is impossible in corpus studies in general, insofar as speaker information is coded and a sufficient amount of data is available.

168 variability is also active at the speaker level. The third goal is to see whether the frequency ef-
169 fect extends to novel compounds composed of existing morphemes of varying frequencies, or
170 whether it is limited to whole forms with which the speakers might plausibly have stored in their
171 mental lexicon.

172 The status of novel compounds is of great relevance for distinguishing between phonolog-
173 ical theories of frequency effects: The USELISTED theory of Zuraw (2000) holds that effects of
174 frequency in phonology, such as those discussed in Section 1, can be explained as competition
175 between two routes of processing—whole-word retrieval, or in-the-moment grammatical assem-
176 bly. The data from existing compounds are compatible with this architecture, but also one where
177 the working of the phonological grammar itself is influenced by frequency. Novel compounds are
178 an important testing ground for theories that put frequency in the lexicon, and finding an effect
179 of N2 frequency in modulating nasalization in these forms would indicate that the phonological
180 grammar itself is sensitive to the lexical frequency of the items in manipulates, as Coetzee and
181 Kawahara (2013) and Coetzee (2016) have suggested.

182 Finally, the experiment included a priming manipulation, designed to probe whether other
183 characteristics of the lexical entry other than frequency might impact their phonological behavior.
184 Breiss (2021) found in several experiments on English derivational morphology that priming non-
185 local paradigm members influenced the way that novel derivatives were formed; for example, a
186 novel derivative in *-able* based on the stem *lábor* was more likely to be pronounced *labórabable*,
187 with stress on the second syllable, when the form *labórious*, with matching stress placement, was
188 primed. Breiss (2021) implemented the manipulation by performing a vocabulary check with
189 a random subset of non-local paradigm members (like *labórious*) before the main task of *-able*-
190 formation, with the rest after; thus, for any individual, half of the items were primed, and the
191 others were not. In this experiment, all but one of the participants completed two repetitions of
192 the experiment, so on each of two runs the participants saw one half of the items primed.

193 3.2 Methods

194 Supplementary material of this experiment, as well as that of Experiment 2, can be accessed at
195 https://osf.io/avnppw/?view_only=cd2afdcc183f4de3ac1261b4af66f08d.

196 3.2.1 Stimuli

197 Stimuli for the experiment were 301 compounds, 81 of which were existing forms (largely drawn
198 from the corpus used in Breiss et al. (2021)), and the rest of which were novel. The complete list
199 of stimuli is available at in the supplementary material. The existing compounds were selected
200 to represent a range of nasalization probabilities, based on the frequency of the whole compound
201 and that of the N2. Due to the challenges in controlling length and morphological composition,

202 there was variation in these aspects among the existing compounds. Specifically, their lengths
203 ranged from 2 to 8 mora in total, and certain compound members were multimorphemic (e.g.,
204 /kagaku/ + /giʒutsu/ 'science and technology'; both N1 and N2 are bimorphemic, i.e., /ka + gaku/
205 "science", /gi + ʒutsu/) "technology"). Existing compounds were all formed with /g/-initial N2s.
206 The novel compounds were formed by combining six bimoraic N1s (e.g., /d̥ʒuu/ "heavy, multiple",
207 /tei/ "low") with 30 bimoraic N2s (e.g., /gin/ "silver", /gʒaku/ "reverse"), with the N2s selected to be
208 of varying frequencies. Examples of novel compounds are /d̥ʒuugin/ (/d̥ʒuu/ "heavy, multiple" +
209 /gin/ "silver") and /toogan/ (/too/ "Chinese" + /gan/ "wild goose"). Due to the difficulty of finding
210 enough monomorphemic N2s, 10 out of the 30 N2s were bimorphemic (e.g., /ge + ta/ "wooden
211 clogs", /go + ma/ "sesame"). The study also included 40 novel compounds whose N2 was /k/-
212 initial, in order to examine the synchronic status of the opaque interaction of nasalization with
213 Rendaku (for which, see the extensive discussion in Ito and Mester 1996, 2003). These results are
214 not reported here, and are thus not discussed further; the data reported for novel compounds in
215 this paper is the result of 180 distinct novel compounds with a /g/-initial N2.

216 3.2.2 Participants

217 We recruited 20 speakers of the Tōhoku dialect of Japanese by word of mouth and snowball
218 sampling to participate in the experiment. We chose to examine the Tōhoku dialect because most
219 (if not all) of the speakers of the Yamanote dialect documented in the literature and reflected in
220 the NHK Pronunciation and Accent dictionary are no longer living, or were judged unlikely to
221 be able to participate in an online experiment. The Tōhoku dialect, spoken in northern parts
222 of the main land of Japan is also phonologically conservative, and has been documented as also
223 exhibiting the voiced velar nasalization alteration. We should nevertheless bear in mind that the
224 Tōhoku dialect is a dialect that is different from the Yamanote dialect that was analyzed by Breiss
225 et al. (2021).

226 All participants completed a short dialect questionnaire, which used existing monomorphemes
227 to determine whether the speaker enforced the complementary distribution of [g] and [ŋ] – the
228 phonotactic restriction, which drives the alternation in compounds. If the speaker did not, they
229 were not invited to continue to the experimental task. Of the 20 interviewed, eight passed the
230 dialect questionnaire.⁴ As two of the three goals of the experiment (see Section 3.1) address ques-
231 tions at the level of the individual speaker's grammar, all but one person (Speaker 7) participated
232 in the experiment in two separate rounds, each separated by a period of a few weeks to sev-
233 eral months. The two different rounds counterbalanced stimulus randomization orders, and also
234 which N2s were primed (on which, see section 3.2.3 immediately below). The participants were

⁴Many speakers, particularly younger ones, have lost this pattern of nasalization due to the influence of the "standard" Tokyo dialect, which has also lost the nasalization pattern, retaining [g] for /g/ in all contexts.

235 paid approximately 20 USD per experimental session.

236 3.2.3 Procedure

237 The format of the experiment was an elicited production task; participants were presented with
238 a series of forms via PowerPoint presentation, and were asked to produce them aloud. All of the
239 experimental sessions were carried out over Zoom by the second author, and were recorded for
240 posterity. As the participant carried out the production task, the second author coded based on
241 auditory impression whether [ŋ] or [g] was produced on a given trial. The decision was usually
242 clear-cut for the experimenter, who is a phonetically-trained linguist and also a native speaker
243 of Japanese. In cases of uncertainty, tokens were coded based on visual inspection of the spec-
244 trogram: tokens involving clear striations for bursts were coded as [g] while ones without them
245 were coded as [ŋ]. The present study does not address the question of phonetic gradience in the
246 implementation of nasalization, for to the best of our knowledge, no studies have demonstrated
247 a gradient nature of the nasalization process; future analysis of the recorded data, however, may
248 yield insight into these questions. Here, we follow the practice in the sociolinguistic (Hibiya,
249 1995) and phonological literature (cited above) and treat nasalization as binary, reflecting the
250 intended outcome of the speaker's grammar.

251 Each recording session proceeded in the following way: first, participants were given the di-
252 alect questionnaire; if they passed, they proceeded to the main task. In the main task, participants
253 completed a preliminary vocabulary familiarity survey before producing compound forms, and
254 a post-hoc vocabulary familiarity survey after producing all the compounds. In each vocabulary
255 survey, participants were asked to produce one of the existing compounds or one of the N2s out
256 loud, and indicate how familiar they were with the word on a 5-point Likert scale (5 = *extremely*
257 *familiar*, 1 = *I don't know this word*).

258 As noted in section 3.1, we took advantage of the two experimental sessions with each partic-
259 ipant to examine what effect priming an N2 might have on whether it exhibits nasalization when
260 produced in a compound. All of the existing compounds included as stimuli were always in the
261 post-hoc vocabulary survey, while the N2s were distributed around the compound-production
262 task in a way that each participant saw one set of N2s before the compound-production task, and
263 one after. In participants who participated in the experiment twice (all but one of them), the N2s
264 primed were varied between sessions.

265 In the compound production task itself, which is of primary interest, participants were asked
266 to simply read aloud compounds followed by a semantically neutral, short sentence-ending par-
267 ticle. These compounds, a mix of novel and existing forms, were presented in a random order.
268 All compounds were presented in *kanji* characters, which do not distinguish between [g] and
269 [ŋ]. Thus we can infer that whatever form the speakers produced ([g]-ful or [ŋ]-ful) is minimally

270 biased by the experimental setup.

271 3.3 Statistical analyses

272 After the data collection was complete, each compound member was classified as whether it was
273 known to the speaker (familiarity score > 1) or not (familiarity score = 1). Then, for each speaker,
274 compounds with unknown N2s were excluded. This allows us to make inferences about the
275 phonological grammar at the level of the speaker, rather than simply assuming that all speakers
276 know all words. All data and scripts are available in the supplementary materials of this paper,
277 available at the osf repository (see above for the link) .

278 Statistical analyses were carried out using Bayesian mixed effects logistic regression models
279 implemented in the *brms* package (Bürkner, 2017) using the R programming environment (R Core
280 Team, 2021). There are several advantages of Bayesian models as opposed to frequentist (non-
281 Bayesian) ones, which we summarise only briefly here. First, rather than focusing on hypothesis
282 testing, the results of our Bayesian regression models can be interpreted as directly reflecting the
283 distributions of likely values for each parameter. Second, it is known that Bayesian models are
284 more likely to converge than corresponding frequentist linear mixed effects models, the latter of
285 which is especially difficult to achieve convergence with when the model has a complex random
286 structure, i.e., the sort of the model we report below. In a Bayesian model, we formalize our prior
287 knowledge or expectations (if any) about the values of the parameters we are interested in using
288 statistical distributions, and then knit it together with the evidence from the data, producing a
289 posterior distribution of values for our parameters of interest that are a compromise between our
290 priors and our data. This posterior distribution is the object which we mine for analytic insights.
291 For more comprehensive tutorial introductions to Bayesian data analysis applied to linguistic
292 and related subject material, see Kruschke (2014), Vasisht et al. (2018); for a primer on the *brms*
293 package specifically in a linguistic context see Nalborczyk et al. (2019).

294 In this paper, we report two common metrics of the posterior distribution for model param-
295 eters of interest: the median and 95% Credible Interval (CI) which is presented as a bracketed
296 range, and the probability of direction, noted $P(|\hat{\beta}| > 0)$. The first measure indicates the median
297 posterior value of the parameter, and the range which encompasses the central 95% of likely val-
298 ues. The second measure can be taken as a way of assessing the amount of evidence we have
299 in favor of any effect in the direction of the parameter coefficient, regardless of magnitude; this
300 ranges from 0.5 (equal evidence for an effect in the direction of the parameter as in the opposite
301 direction) to 1 (very strong evidence in favor of an effect in the direction of the parameter value).

302 In terms of model structure, each model used as its dependent variable the realization of the
303 initial segment of N2 ([g] or [ŋ]), contained fixed effects specified below, and random intercepts
304 for speaker and compound, with random slopes of all fixed effects by speaker and a random slope

305 of priming (primed or not primed) by compound. The models used Normal (0,1) priors on the
306 intercept and coefficients; sensitivity analyses (Roos et al., 2015) revealed no meaningful changes
307 in inference were associated with a range of prior values, indicating that the data we collected
308 were sufficiently informative that our prior beliefs about likely parameter values mattered only
309 nominally; see the supplementary materials for details.

310 We pause here to draw attention to the fact that out of a desire to have enough types of real and
311 novel compounds, spanning a range of frequencies, we were unable to make the existing com-
312 pounds uniform in size, and neither the existing nor novel compounds uniform in morphemic
313 composition. Because we have no reason to believe these factors to be causally related to the
314 propensity to undergo nasalization, and on the basis of the second two authors' native speaker
315 intuition that the bimorphemic compound members were much more salient as whole words,
316 rather than compositionally-formed parts, we do not consider these as theoretical quantities of
317 interest in our statistical or grammatical analysis. We expect the random intercept for compound
318 included in all of the statistical models we fit to absorb any idiosyncrasy attributable to mor-
319 phemic composition or length to be absorbed by this term, treated as item-level quirks that need
320 to hold from sample to sample, in the same way that idiosyncratic participant-level variation is
321 absorbed by the random intercept for participant. Readers interested in investigating the causal
322 link between nasalization and these other factors for themselves may access the raw data in the
323 supplementary materials.

324 **3.4 Results**

325 In what follows, we first visualize and qualitatively discuss the results of the experiment, then
326 perform parameter estimation using a Bayesian model to conform the statistical reliability of our
327 observations.

328 **3.4.1 Existing compounds**

329 We first examine existing compounds with /g/-initial N2s, and ask whether the token frequency
330 of the compound or the N2 explains variation in nasalization; this is plotted in Figure 4. Note that
331 while Breiss et al. (2021) considered *relative* frequency of the N2, here we consider the frequency
332 of the N2 and compound in their own right. Finally, we use the natural logarithm of the token
333 frequency, rather than its raw value, as is standard practice.

334 We find that the frequency effect is robust, bearing out the spirit of the effect seen in the
335 lexical analyses reported in Breiss et al. (2021). As N2 frequency rises relative to a fixed value
336 of compound frequency, the probability of an individual compound exhibiting nasalization drops
337 (Figure 4, right facet); holding N2 frequency steady while increasing the frequency of a compound
338 from low to high also increases the probability of nasalization (Figure 4, left facet).

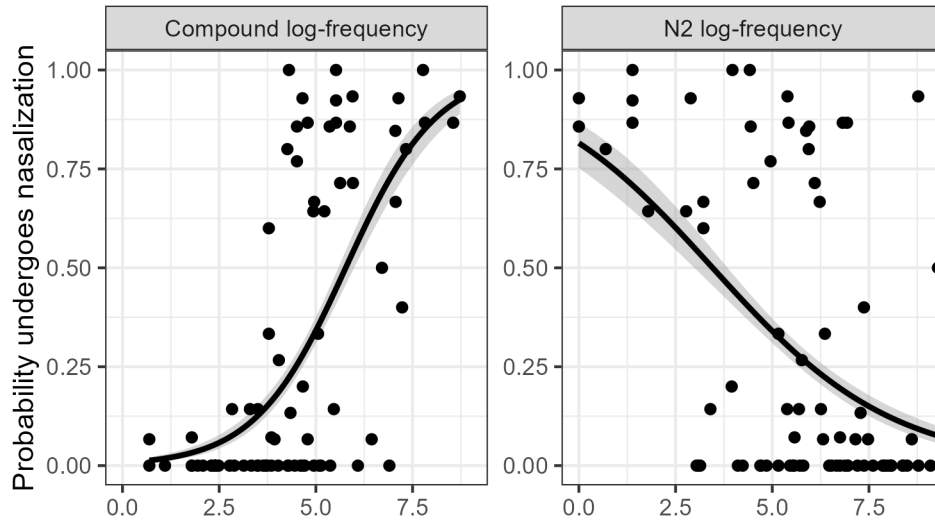


Figure 4: Probability of nasalization (vertical axis) plotted against compound log-frequency (left facet) and N2 log-frequency (right facet), with binomial smooths.

339 Having found that paradigm uniformity is conditioned by frequency in existing compounds
 340 at the group level, we now examine whether the conditioning holds at the level of the individual
 341 grammar, plotting each participant in their own row, in Figure 5.

342 On visual inspection, it seems clear that the effect does exist at the individual level, but may
 343 differ in strength between speakers. We return to this question with a quantitative eye when we
 344 discuss the statistical model fit to this data.

345 Before doing so, however, we report two null results in the set of existing compounds that
 346 we had expected to find based on the literature: that of priming the N2, and of OCP[nasal]. We
 347 had anticipated that priming the N2 would impact the likelihood of compounds with primed N2s
 348 to undergo nasalization, based on the findings and rationale of Breiss (2021), described above in
 349 section 3.2.3. However, priming seemed to have no meaningful effect on rates of nasalization (left
 350 plot of Figure 6). Thus, we conclude that the experimental manipulation (placement of the N2
 351 in the vocabulary check sequence) failed to influence the salience of the N2 in the lexicon of the
 352 participant in such a way for it to be experimentally detectable in their behavior on the production
 353 task. Though further research is needed to confirm, we suspect that the lack of priming in this
 354 study was because the manipulation tried to target too many N2s at once, leading to a lack of
 355 concentrated activation on any particular item relative to the rest. This post-hoc account predicts
 356 that studies that try to prime relatively fewer items (as was the case with Breiss 2021) should
 357 have a greater chance of estimating the effect of priming on the phonological grammar, but more
 358 targeted work on this topic is needed to better support this hypothesis.

359 Second, we expected, based on the findings of Breiss et al. (2021) in Japanese compounds and

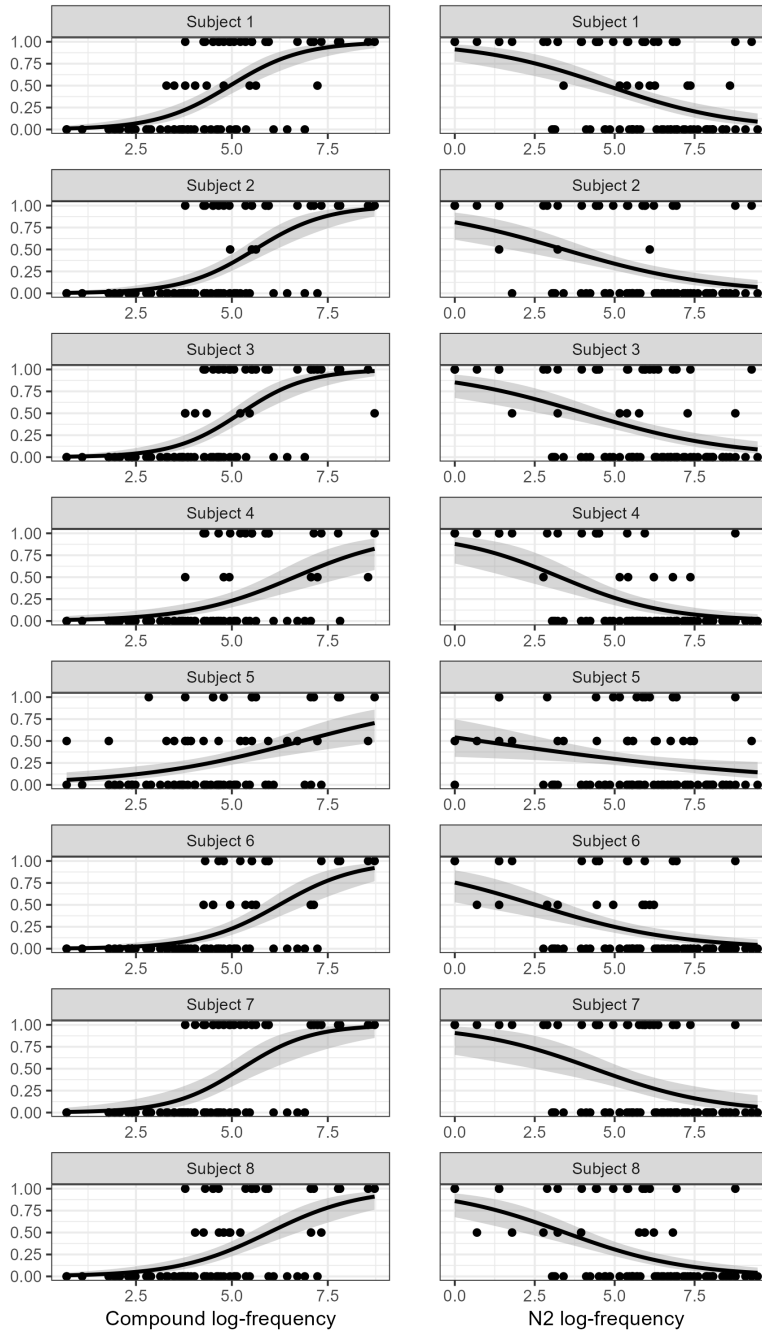


Figure 5: Probability of nasalization (vertical axis) plotted against compound log-frequency (left facet) and N2 log-frequency (right facet) for each individual speaker (row), with binomial smooths.

360 more broadly in Japanese phonology (Kawahara et al., 2006), to find a decrease in nasalization
 361 in compounds whose N1s were nasal-final, so as to avoid creating a sequence of two nasals. We
 362 found the opposite trend in the experimental data, as shown in the right panel of Figure 6, but
 363 only superficially. Although the left bar is higher than the right bar, the uncertainty about this

364 measure is also much larger, likely stemming from the relatively few compounds ($n = 16$) that
 365 have nasal-final N1s, and so the statistical model we fit does not suggest that the visual trend is to
 366 be trusted. Based on the conflicting evidence in the literature and this paper, we make no strong
 367 conclusions about the interaction of OCP[nasal] and nasalization in Japanese, and await future
 368 more targeted experiments that address this question directly.

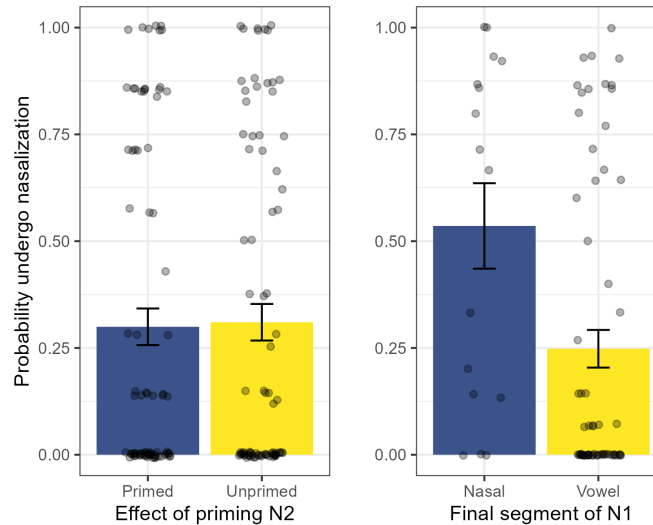


Figure 6: Probability of nasalization with standard error (vertical axis) plotted against priming of N2 (left) and final segment of N1 (right).

369 Table 1 presents the results of a Bayesian mixed-effects logistic regression model fit to deter-
 370 mine the statistical robustness of the data patterns just reviewed. The model structure and random
 371 effects was as described in section 3.3, and included fixed effects of the scaled log-frequency of
 372 the N2 and compound, whether N2 was primed, the interaction of priming with frequency of N2,
 373 and the nasality of the final segment of N1.

374 We note that, relative to the intercept, a one-unit increase in compound log-frequency strongly
 375 increases the log-odds of the compound undergoing nasalization, and a one-unit increase in N2
 376 log-frequency decreases the log-odds of nasalization — we judge this by the fact that the central
 377 95% of credible values for the two frequency coefficients exclude zero. For all other fixed effects,
 378 the 95% CI does include zero, so we are less confident in attributing a meaningful effect on the
 379 data to these factors.

380 We estimate the speaker-specific parameter value for both compound and N2 frequency by
 381 examining samples extracted from the model; these are summarised in Table 2 using the same
 382 metrics as for the model in Table 1. For all participants, an increase in N2 frequency was associ-
 383 ated with a decrease in nasalization with greater than 99% probability. For compound frequency,
 384 an increase in frequency was associated with a decrease in nasalization with greater than 90% con-
 385 fidence for six of eight speakers; for Speaker 4 the effect was less certain (80%), and only Speaker

<i>Parameter</i>	<i>Median</i>	<i>95% CI</i>	<i>P</i> ($ \hat{\beta} > 0$)
Intercept:			
N1 final segment = <i>vowel</i>			
N2 primed = <i>no</i>			
Compound scaled log freq. = <i>mean</i>			
N2 scaled log freq. = <i>mean</i>	-0.92	[-2.34, 0.53]	
N1 final segment = <i>nasal</i>	0.61	[-0.72, 1.94]	0.82
N2 primed = <i>yes</i>	-0.30	[-0.31, 1.91]	0.84
Compound log freq. (<i>one unit increase</i>)	2.40	[0.84, 3.55]	0.99
N2 log freq. (<i>one unit increase</i>)	-1.42	[-2.20, -0.53]	0.99
N2 log freq. \times N2 primed = <i>yes</i>	-0.35	[-0.91, 0.21]	0.91

Table 1: Model of existing compounds with free N2s. Coefficients are in log-odds, with positive signs indicating an increase in probability of nasalization relative to the intercept.

386 5 truly exhibited no evidence for compound frequency influencing nasalization rate (though the
387 same speaker exhibited a strong influence of N2 frequency).

388 Based on this evidence, we think it is reasonable to impute the frequency effect in existing
389 compounds to the level of the individual grammar, though the factors influencing individual-level
390 variation in effect size remain for future research.

391 3.4.2 Novel compounds

392 Turning to novel compounds, we find that the frequency effect holds here as well, though with a
393 smaller magnitude. Figure 7 plots only N2 frequency; since the the compound is entirely novel, its
394 frequency is naturally zero. The downward-sloping smooth qualitatively matches the one found
395 in the right panels of Figures 4 and 5.

396 Breaking this result out by individual in Figure 8, we find that visually there appears to be a
397 wide range of variation in the strength of the effect across participants, though all but one go in
398 the expected direction. We return to by-subject estimates from a fitted model below.

399 Finally, consistent with the null effect observed in the existing compounds, we found no
400 strong evidence that priming the N2 influenced nasalization application in novel compounds;
401 this is shown in Figure 9. Since none of the six N1s we selected for constructing the novel com-
402 pounds were nasal-final, we were not able to evaluate the effect of OCP[nasal] in this subset of
403 the data.

404 The results of a Bayesian logistic regression model fit to the data for compounds with novel
405 /g/-initial N2s are reported in Table 3.

406 Consistent with the existing compounds, we find a strong effect of N2 log-frequency, with

	Parameter	Median and 95% CI	$P(\hat{\beta} > 0)$
<i>Speaker 1</i>	Compound log freq.	3.99 [2.02, 6.04]	≈ 1
	N2 log freq.	-1.35 [-2.65, -0.04]	0.98
<i>Speaker 2</i>	Compound log freq.	2.53 [0.78, 4.33]	0.99
	N2 log freq.	-2.56 [-3.93, -1.24]	≈ 1
<i>Speaker 3</i>	Compound log freq.	3.57 [1.69, 5.56]	0.99
	N2 log freq.	-1.64 [-2.93, -0.30]	0.99
<i>Speaker 4</i>	Compound log freq.	0.74 [-0.96, 2.54]	0.80
	N2 log freq.	-4.78 [-6.77, -3.03]	≈ 1
<i>Speaker 5</i>	Compound log freq.	0.04 [-1.40, 1.45]	.52
	N2 log freq.	-1.87 [-3.06, -0.71]	.99
<i>Speaker 6</i>	Compound log freq.	1.52 [-0.49, 2.85]	.91
	N2 log freq.	-3.99 [-5.53, -2.46]	≈ 1
<i>Speaker 7</i>	Compound log freq.	3.77 [1.44, 6.46]	.99
	N2 log freq.	-2.07 [-3.75, -4.94]	.99
<i>Speaker 8</i>	Compound log freq.	1.56 [-0.07, 3.13]	.97
	N2 log freq.	-3.22 [-4.67, -1.89]	≈ 1

Table 2: Summaries of individual-level estimates of the effect of the two frequency parameters derived from the model in Table 1.

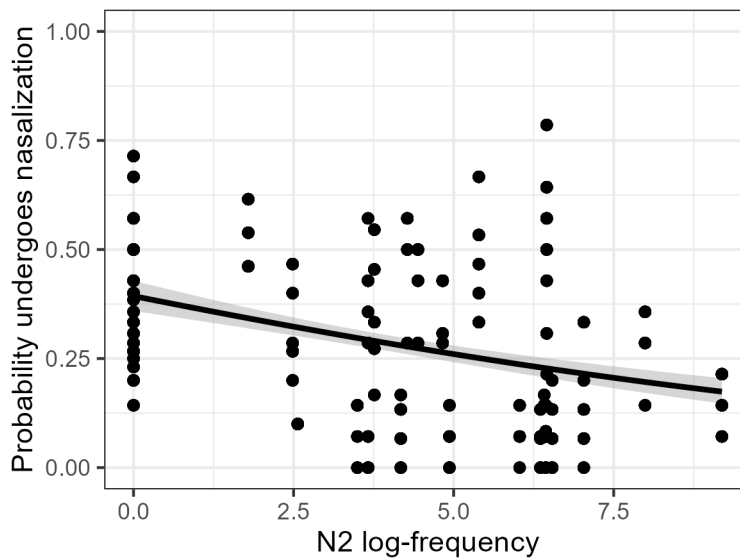


Figure 7: The probability of undergoing nasalization, plotted against N2 log-frequency (novel compounds), with a binomial smooth to aid readability.

407 greater values inhibiting nasalization. None of the other main effects were statistically reliable.
 408 At the individual level, we find strong evidence for a frequency effect of N2 in all individuals;
 409 in all but one, the effect is as expected, with higher frequency N2s forming compounds that

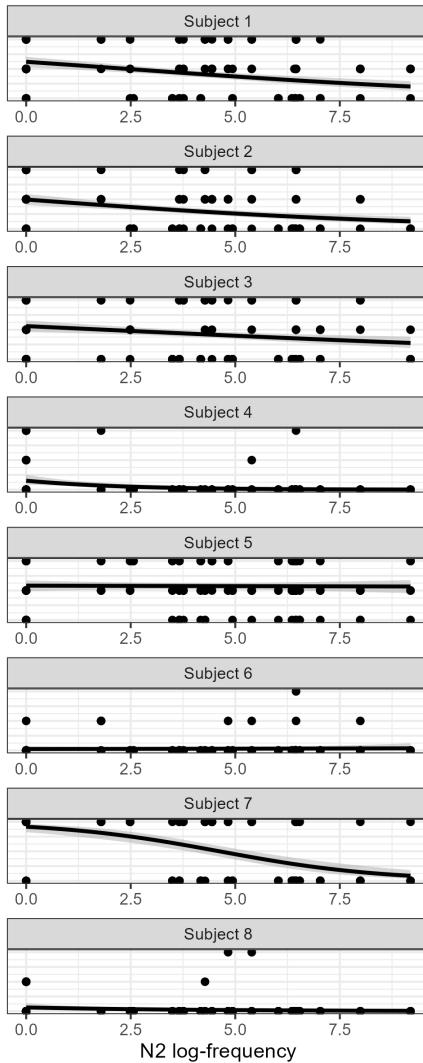


Figure 8: Probability of nasalization (vertical axis) plotted against N2 log-frequency (horizontal axis) for each individual (row), with binomial smooths.

410 are less likely to exhibit nasalization. In one individual, Speaker 5, however, the effect is in the
 411 opposite direction; this is unexpected, and further puzzling because the same speaker exhibits a
 412 robust frequency effect of N2 in the expected direction in existing compounds (though no strong
 413 evidence for an effect of compound frequency in that data, interestingly). We have no explanation
 414 for this pattern, other than to note that the effect holds in all other participants; future work
 415 is needed to understand the factors that might yield different effects of frequency in different
 416 individuals.

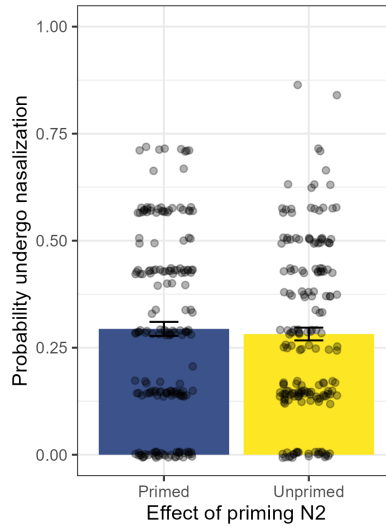


Figure 9: The probability of undergoing nasalization with standard error (vertical axis) based on whether the N2 was primed (horizontal axis) in novel compounds.

<i>Parameter</i>	<i>Median</i>	<i>95% CI</i>	$P(\hat{\beta} > 0)$
Intercept:			
N2 primed = <i>no</i>			
N2 scaled log freq. = <i>mean</i>	-1.92	[-3.68, -1.11]	
N2 primed = <i>yes</i>	0.04	[-0.43, 0.53]	0.57
N2 log freq. (<i>one unit increase</i>)	-0.50	[-0.96, -0.02]	0.98
N2 log freq. \times N2 primed = <i>yes</i>	-0.03	[-0.38, 0.33]	0.58

Table 3: Model of novel compounds with free N2s. Coefficients are in log-odds, with positive signs indicating an increase in probability of nasalization relative to the intercept.

	Median and 95% CI for effect of N2 log-frequency	Probability of direction
<i>Speaker 1</i>	-1.59 [-2.30, -0.93]	≈ 1
<i>Speaker 2</i>	-1.98 [-2.65, -1.31]	≈ 1
<i>Speaker 3</i>	-0.98 [-1.55, -0.35]	0.99
<i>Speaker 4</i>	-5.64 [-7.46, -4.16]	≈ 1
<i>Speaker 5</i>	0.69 [0.07, 1.36]	0.99
<i>Speaker 6</i>	-4.39 [-5.66, -3.27]	≈ 1
<i>Speaker 7</i>	-1.73 [-2.72, -0.68]	0.99
<i>Subject 8</i>	-5.48 [-7.05, -4.08]	≈ 1

Table 4: Summaries of individual-level estimates of the effect of the N2 frequency parameter derived from the model in Table 3.

417 **3.4.3 Within-speaker consistency**

418 Finally, we examined within-speaker consistency in the frequency effect, comparing existing
 419 compounds to novel compounds. If the frequency effect is active at the level of the individual
 420 speaker, it is reasonable to assume that some speakers might be more sensitive or affected by fre-
 421 quency than others, and that this effect impacts the treatment of existing and novel compounds
 422 by the phonological grammar. Figure 10 plots the coefficient of N2 frequency in novel compounds
 423 (vertical axis) against that of N2 frequency in existing compounds (horizontal axis); points repre-
 424 sent the median value of the posterior, and the error bars encompass the 95% CI. A linear smooth
 425 is added for readability, and a red dashed line with slope 1 is provided for comparison.

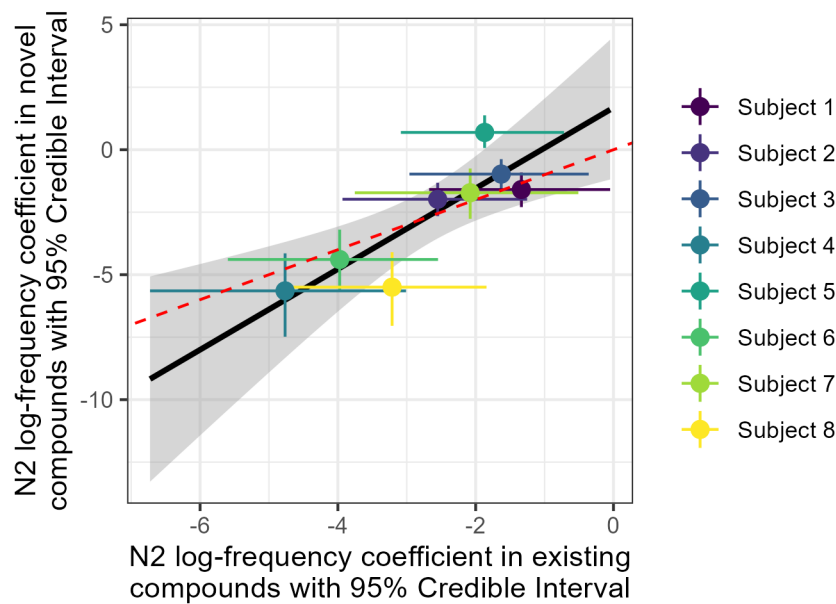


Figure 10: The coefficient of N2 log-frequency in novel compounds, derived from the model in Table 1, is plotted on the horizontal axis, and the coefficient for N2 log-frequency in existing compounds, derived from the model summarised in Table 3, is plotted on the vertical axis. Points represent median values of the posterior, and ranges encompass 95% CI, colors represent speakers, and a linear smooth has been added for readability, with the line of slope 1 intersecting the origin in dotted red.

426 While it is clear that individual speakers exhibit variation in how strongly they are affected
 427 by lexical frequency (cf. the differences of magnitude of the relevant median coefficient values
 428 in Tables 1 and 3), what is important to note here is that individuals are self-consistent in their
 429 variation — speakers who are more sensitive to N2 log-frequency in existing compounds are also
 430 more sensitive in novel compounds, and similarly for those that are less sensitive (Pearson’s $\rho =$
 431 0.86 , $p = 0.006$). Further, we can qualitatively observe that the effect of frequency is of roughly
 432 the same magnitude across both compound types, corresponding to the overlap between the grey

433 smooth’s uncertainty and the red dashed line of identity in Figure 10. This suggests that how-
434 ever token frequency influences the phonological grammar, it does so indiscriminately, ignoring
435 the phonological or morphological nature of the items themselves (here, existing vs. novel com-
436 pounds).

437 In summary, Experiment 1 demonstrates that the conditioning effect of both compound fre-
438 quency and N2 frequency holds in productions of existing compounds, as suggested by the corpus
439 analysis of Breiss et al. (2021). It also demonstrates that the frequency-conditioning holds at the
440 level of the individual speaker, precluding an explanation of the variable pattern in terms of av-
441 eraging across individual categorical grammars. Finally, and most importantly for the construc-
442 tion of phonological theories of lexicon-grammar interaction, it confirms that the frequency of
443 N2 plays a role even in entirely novel compounds, indicating that the grammar must be sensitive
444 to the lexical frequency of morphemes it manipulates.

445 4 Experiment 2

446 Experiment 2 complements the coverage of Experiment 1 by systematically varying the length
447 of the compound (two, three, or four mora), and the frequency of N2. It also more carefully con-
448 trols for morphemic composition, which was somewhat compromised in Experiment 1 in favor
449 of having a greater range of frequencies and a sufficient number of stimuli. Recall that the goal
450 of the length manipulation in Experiment 2 was testing whether the corpus pattern that sug-
451 gested a relationship between overall compound size and propensity to nasalize, a possible case
452 of phonology “counting”. Although Jiang (2023b) carried out a similar test of the relationship be-
453 tween nasalization and compound size and found that it did not generalize to novel compounds,
454 his study was not restricted to only speakers of a phonologically-conservative dialect which pre-
455 serves the nasalization pattern. Therefore we consider it important to double check this finding,
456 especially given that it touches on a core question of phonological theory. Finally, since the stim-
457 uli in Experiment 2 are all novel compounds, it serves as a replication of the frequency effect in
458 novel compounds from Experiment 1.

459 4.1 Methods

460 4.1.1 Stimuli

461 Six N1s were selected, three monomoraic and three bimoraic (e.g., [ta] “many”, [tai] “pair”), and
462 were fully crossed with 28 N2s, roughly balanced between monomoraic ($n = 12$; e.g., [go] “Go”,
463 [ga] “moth”) and bimoraic ($n = 16$; e.g., /gin/ “silver”, /gjaku/ “reverse”). All compound members
464 were monomorphemic. Within each N2 length, stimuli sampled a wide range of log frequencies

465 (monomoraic: 1.39-7.40, mean 4.61; bimoraic: 3.56-8.0, mean 5.51), and did not differ significantly
466 in overall frequency via a two-sided t -test ($t = -1.41$, $df = 18.98$, $p = 0.17$). This yielded 168 novel
467 compounds for analysis.

468 We also included 32 compounds where the N2 was /k/-initial, as part of a separate exper-
469 imental condition designed to test the interaction of Rendaku and nasalization. The research
470 questions at stake in this condition are disjoint from those pursued in this paper, so these stimuli
471 are excluded from our analysis and not discussed further.

472 4.1.2 Participants

473 A total of 13 participants were recruited to participate in Experiment 2 via word of mouth and
474 snowball sampling. As in Experiment 1, each participant was screened using the dialect ques-
475 tionnaire to ensure that they spoke a dialect where [g] and [ŋ] are allophonically distributed in
476 monomorphemic words. Of the 13 interviewed, 12 passed the dialect questionnaire (see section
477 3.2.2). Three participants also participated in Experiment 1, while the other nine did not.⁵ Unlike
478 Experiment 1, participants only participated in this experiment once, due to time and resource
479 constraints.

480 4.1.3 Procedure

481 Procedure for Experiment 2 was identical to Experiment 1, as was data processing and statistical
482 analysis.

483 4.2 Results

484 We begin with the central question of Experiment 2: does compound length in mora – above and
485 beyond other factors, like frequency – influence how often speakers apply nasalization in novel
486 compounds? Figure 11 plots the mean nasalization rate by compound length and N2 frequency.

487 We find that shorter compounds appear to undergo nasalization more often than longer com-
488 pounds, suggesting that the effect of global prosodic length seen in Figure 3 may be active in the
489 grammar of Tōhoku Japanese speakers; the N2 frequency effect in novel compounds observed in
490 Experiment 1 seems to replicate in Experiment 2.

491 To assess the credibility of the visual trends observed in Figure 11, we fit a Bayesian mixed-
492 effects logistic regression model to the experimental results, which is given in Table 5 below.
493 The model structure followed the outline in section 3.3, and had a three-level ordered factor of

⁵While having speakers in common between the two experiments may not be ideal, it was quite difficult to find speakers who both spoke the conservative dialect and also were comfortable with participating in an experiment using Zoom. The two experiments were conducted over the span of three years and therefore we think it unlikely that detailed memory of the first experiment influenced their performance in the second.

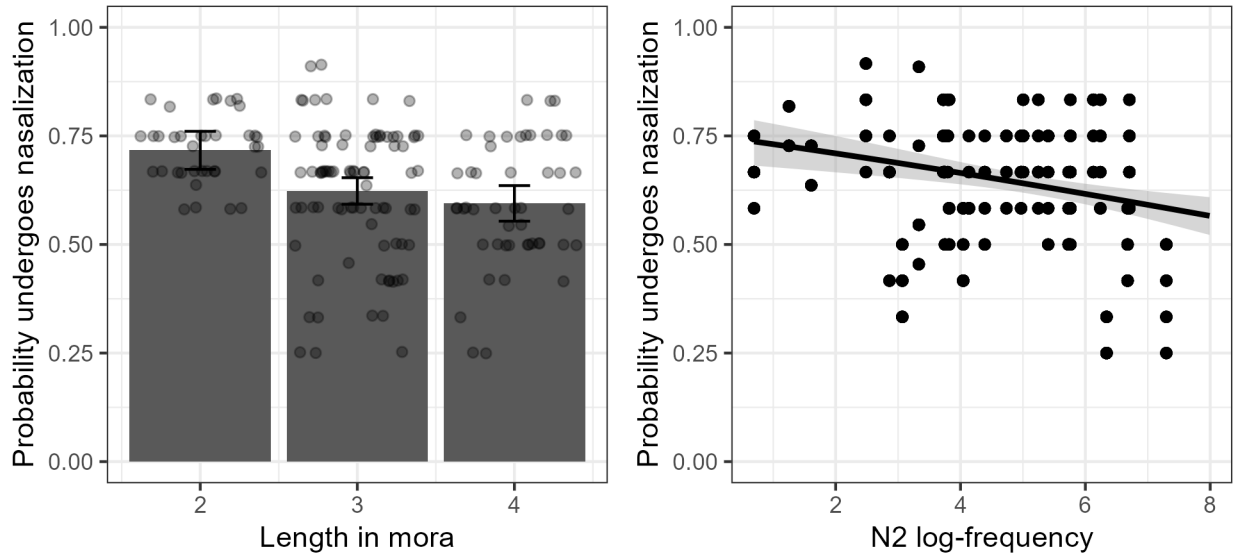


Figure 11: Probability of nasalization plus standard error (vertical axis) plotted against the length of the compound in mora (horizontal axis, left) and N2 log-frequency (horizontal axis, right), with a binomial smooth.

494 compound length in mora (levels 2, 3, and 4), N2 log-frequency, whether N2 was primed, and the
 495 interaction between priming and N2 log-frequency. Since all N1s in Experiment 2 were vowel-
 496 final, we did not assess possible avoidance of adjacent nasals across the N1-N2 boundary.

497 The negative coefficient for N2 log-frequency indicates that there is substantial statistical sup-
 498 port for the visual impression that compounds with higher-frequency N2s undergo nasalization
 499 less often. Post-hoc pairwise comparisons between levels of *Mora in compound* using the *em-*
 500 *means* package (Lenth et al., 2019) revealed that three-mora compounds underwent nasalization
 501 at a credibly lower rate than two-mora compounds, following the visual intuition in Figure 11, but
 502 that the difference in nasalization rate for three- and four-mora compounds was of a smaller mag-
 503 nitude, and had a posterior distribution of credible values that overlapped zero, though around
 504 80% of the credible values are compatible with a small positive effect ($\hat{\beta} = 0.28$, 95% CI [-0.44, 1.00],
 505 $p(|\hat{\beta}| > 0) = 0.79$). Finally, as in Experiment 1, neither priming nor its interaction with N2 log-
 506 frequency meaningfully predicted nasalization in participant responses (not plotted for space). In
 507 sum, Experiment 2 replicates the finding from Experiment 1 that the phonological treatment of
 508 even novel compounds are subject to the lexical frequency of their N2s, and also provides partial
 509 support for the global prosodic length effect observed in the corpus, a topic to which we now
 510 turn in more depth.

<i>Parameter</i>	<i>Median</i>	<i>95% CI</i>	$P(\hat{\beta} > 0)$
Intercept:			
Mora in compound = 2			
N2 primed = <i>no</i>			
N2 scaled log freq. = <i>mean</i>	1.65	[-0.72, 4.08]	
Mora in compound = 3	-0.78	[-1.43, -0.10]	0.98
Mora in compound = 4	-1.07	[-1.82, -0.32]	≈ 1
N2 primed = <i>yes</i>	0.15	[-0.39, 0.67]	0.74
N2 log freq. (<i>one unit increase</i>)	-0.41	[-0.83, 0.01]	0.97
N2 log freq. \times N2 primed = <i>yes</i>	-0.13	[-0.59, 0.31]	0.71

Table 5: Model of novel compounds with free N2s in Experiment 2. Coefficients are in log-odds, with positive signs indicating an increase in probability of nasalization relative to the intercept.

4.3 Discussion

Since the planned analysis of overall compound length revealed only a partial effect, we carried out an exploratory analysis on the compounds to better understand the nature of whatever length effect might exist.

In contrast to the gentle downward trend in nasalization for longer compounds seen in the aggregated data (Figure 11), when we examine the data at the level of the compound’s moraic composition, we find a strikingly different pattern. Rather than being intermediate between two- and four-mora compounds, the two three-mora groups diverge in behavior on the basis of the length of their N2. Compounds with monomoraic N2s pattern with two-mora compounds (which have monomoraic N2s) in having a higher overall nasalization rate, while those that have bimoraic N2s pattern with four-mora compounds (which also have bimoraic N2s). Thus, a more detailed examination of the Experiment 2 results suggests that it is not global compound length, but rather N2 length specifically, that is the dominant determinant of nasalization in the novel compounds. This contradicts the typologically-unusual pattern suggested in Figure 3, suggesting that the speakers of the Tōhoku dialect do not generalize a relationship between global prosodic length and a local segmental alternation (see also similar non-generalization of a similar length-referring pattern in an Artificial Grammar Learning paradigm by Jiang 2023a).

We also observe that the effect of frequency seems to be much more pronounced in compounds with bimoraic N2s (right column of the right graph in Figure 12), compared to those with monomoraic N2s. To assess whether this difference was statistically reliable, we fit a model with structure similar to the one summarised in Table 5, but with fixed effects of N1 mora (1 vs. 2), N2 mora (1 vs. 2), and the interaction of both N1 mora and N2 mora and N2 log-frequency

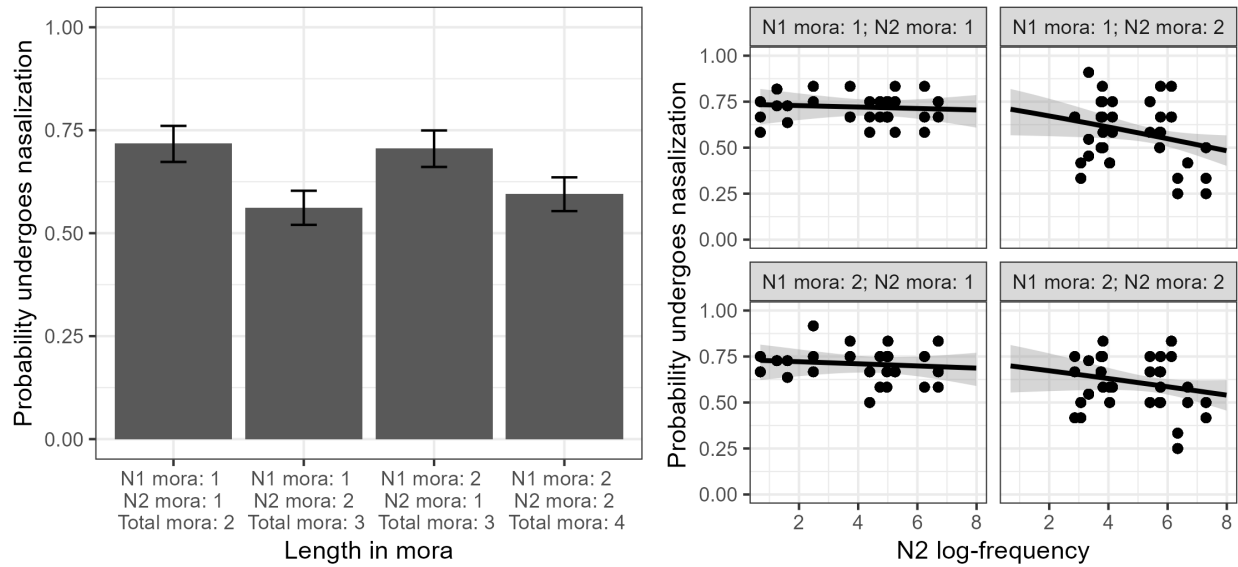


Figure 12: Probability of nasalization plus standard error (vertical axis) plotted against the moraic composition of the compound (horizontal axis, left; panels, right) and N2 log-frequency (horizontal axis within panels, right), with binomial smooths.

533 (thought not the three-way interaction); the model also included the fixed effect of priming and
 534 its interaction with N2 log-frequency. We found that while the effect of N2 mora was strong and
 535 credible ($\hat{\beta} = -1.33$, 95% CI [-2.25, -0.35], $p(|\hat{\beta}| > 0) = 0.99$), the interaction between N2 mora and
 536 N2 log-frequency was not credibly different from zero ($\hat{\beta} = -0.18$, 95% CI [-0.71, 0.35], $p(|\hat{\beta}| > 0) =$
 537 0.76), suggesting that the apparent differences in frequency effect between columns in the right
 538 of Figure 12 are artifacts of conversion from the unbounded space of log-odds which the model
 539 uses, to the bounded interval of probabilities which characterize the data.

540 **4.4 Global prosodic length does not directly influence nasalization**

541 Having observed that N2 length—not overall compound length—is a predictor of nasalization
542 above and beyond frequency, we speculate that the distinct behaviors observed between monomoraic
543 and bimoraic N2s are due to prosodic factors (noting that this is a post-hoc analysis, which
544 needs to be more fully addressed in a new study). Specifically, it may be that nasalization is
545 blocked by prosodic word boundaries, and bimoraic N2s are more likely to form an independent
546 prosodic word than monomoraic N2s for prosodic reasons. We urge caution with this interpre-
547 tation though, since these prosodic domains are typically posited in the research on compound
548 accentuation for Tokyo Japanese, and so because our speculation here is entirely based on this
549 body of work its transportability to a different dialect—Tōhoku Japanese—may be limited.

550 In examining compound accentuation in Tokyo Japanese, Ito & Mester (2018; 2021) propose
551 that monomoraic or bimoraic N2s form a foot, while N2s longer than this project their own
552 prosodic word. This assumption was made to capture the traditional distinction between com-
553 pounds with a “short” N2 and compounds with “long” N2. Although the details of compound
554 accentuation are not our primary concern, in general terms, compounds with a short N2 have
555 compound accent located at the end of N1 (e.g., [tinomi'+go] “suckling child”), while ones with
556 a long N2 either preserve the original accent of N2 (e.g., [hon+ka'igi] “main session”) or place an
557 accent at the beginning of N2 (e.g., [kuchi+ge'nka] “oral quarrel”). The essence of the analysis
558 is that the presence or absence of an accent on N2 depends on whether N2 constitutes an inde-
559 pendent prosodic word. However, this claim is not without exceptions—indeed, Ito and Mester
560 (2018) discuss instances where bimoraic N2s with lexical initial accent retain this initial accent
561 (e.g., [watashi+bu'ne] “ferry boat”), and to account for such cases they assume that these bimoraic
562 N2s exceptionally form a prosodic word.

563 Finally, a question that remains unresolved by this speculation is why the size of N2 should
564 matter, and not the size of N1. A more articulated theory of the relationship between prosodic
565 length, foot structure, and accentuation is required for Tōhoku Japanese is needed, which will
566 then support future experimental work with a larger range of compound member sizes. For the
567 purposes of our discussion here, we can say that Experiment 2 does not support the claim that
568 global prosodic length directly influences nasalization, while noting that the size of the individual
569 compound members does play such a role, though the details of this picture are still quite murky.

570 **5 Implications for phonological theories**

571 To summarise the contributions of the two experiments described above, we found that the vari-
572 ability reported in Ito and Mester (1996, 2003) and Breiss et al. (2021) for the Yamanote dialect
573 is reproduced experimentally with speakers of the Tōhoku dialect in existing compounds (Ex-

574 periment 1) as well as novel ones (Experiments 1 and 2), and showed that this variability holds
575 at the level of each individual speaker. We also investigated a potential case of “phonological
576 counting” where nasalization appeared to be conditioned by overall mora count of the compound
577 (Experiment 2). Instead, we revealed that the nasalization was sensitive only to the size of N2,
578 and speculated what type of metrical structures might — short of a counting-based analysis —
579 account for this pattern.

580 These data pose challenges for existing models of frequency-dependent phonological behav-
581 ior, primarily because we find frequency-dependent behavior exhibited in entirely novel morpho-
582 logical constructions. Many existing theories rely on the speaker maintaining multiple represen-
583 tations at different levels of granularity in their lexicon or long-term memory (here, the com-
584 pound as a unit as well as the N2 as a unit). However, theories that locate frequency in the lexical
585 representation of the larger unit which displays the variation (here, the compound) have difficulty
586 accounting for the probabilistic and frequency-conditioned behavior of morphemes when placed
587 in novel contexts. For example, Zuraw’s USELISTED model (Zuraw, 2000; Zuraw and Peperkamp,
588 2015; Zuraw et al., 2020) proposes that a stored form competes with a grammatically-composed
589 form in speech production, and the relative frequency of the stored form regulates its competi-
590 tiveness. Under this analysis, compounds might be stored with the nasalization process applied
591 (following ideas in Martin 2011), and the nasalized form competes with a grammatically-derived
592 form (cf. the implementation in Zuraw 2007). Depending on the details of the theory, the gram-
593 mar itself might allow probabilistic nasalization in its outputs, but this does not depend on the
594 lexical frequency of the item directly. Thus, any observed frequency-dependent nasalization in
595 existing compounds is the result of greater lexical frequency favoring the direct retrieval of the
596 nasalized form. However, in the case of novel compounds where the frequency of N2 regulates
597 the probability of nasalization, the USELISTED theory has no obvious stored form which competes
598 with the on-line derived one. Therefore, whatever rate of nasalization is set in the grammar must
599 be derived on-line, and cannot relate to frequency.

600 Theories that attribute idiosyncratic behavior to individual morphemes or larger units via
601 constraint indexation or similar mechanism assume that the item which is indexed is itself the
602 exceptional one. In the case of existing items, lexical frequency is often adduced as not only a
603 possible conditioning factor on their variation, but a critical one in both allowing speakers to learn
604 item-specific behavior (Moore-Cantwell and Pater, 2016; Smith and Moore-Cantwell, 2017; Zymet,
605 2018) and to exhibit type-level “frequency matching” of a probabilistic phonological process in
606 novel forms (cf. Zuraw, 2000; Ernestus and Baayen, 2003; Hayes et al., 2009). In the case of entirely
607 novel compounds, however, there is (presumably) no indexed constraint for a never-before-heard
608 item, and so this mechanism also seems insufficiently expressive.

609 Other types of analysis that attach the frequency-conditioning to the item itself (here the N2)

610 fare better with the data from novel compounds. These analyses break down roughly into ones
611 that hold that differing token frequency has representational consequences for the N2 itself, and
612 those that attribute the frequency-conditioning to the constraint violations that the N2 incurs.
613 Examples of analysis in the representation-driven style include those where the first segment
614 of the N2 is represented as a blend of segments based on the frequency of attestation (in free
615 forms or compounds, depending on the implementation) (e.g., Gradient Symbolic Representations
616 (Smolensky and Goldrick, 2016) or Representational Strength Theory (Moore-Cantwell, *ms*)), or
617 where N2 has two stored allomorphs, among which choice is governed by frequency.

618 A second class of theories include those where the lexical frequency of a form may not have
619 direct representational consequences, but appeals (implicitly or explicitly) to the notion of fre-
620 quency as a psycholinguistic quantity which characterizes the accessibility or prominence of
621 lexical representations for the grammar. Phonological models have been proposed that scale the
622 weights (Coetzee and Pater, 2006, 2008; Coetzee, 2016; Coetzee and Kawahara, 2013) or violations
623 (Breiss, 2021) of faithfulness constraints by a quantity that relates to the lexical frequency of the
624 form. Since these types of theories locate the effect of frequency on (or near) the N2 itself, they
625 are at least in a position to address the data from novel compounds.

626 Beyond this, however, it is difficult to say exactly which mechanisms are required to capture
627 the frequency-conditioning in novel forms without an implemented computational model of the
628 experimental data, which we leave for future work. What does seem clear, at least qualitatively,
629 is that theories that enrich the grammar and/or lexicon for only morphologically-complex items
630 (here, just compounds) are likely insufficient.

631 Finally, any theory that works well for the frequency-conditioning data in novel compounds
632 will need to also be able to model the two additional frequency effects at play in existing com-
633 pounds: higher compound frequency (holding N2 frequency constant) increases the likelihood
634 of nasalization, while simultaneously higher N2 frequency (holding compound frequency con-
635 stant) decreases the likelihood of nasalization. A successful model is going to need to be able to
636 capture these effects in the same model of indexation, listing, or representational enrichment –
637 or demonstrate that adopting a hybrid approach achieves empirical coverage that outweighs the
638 penalty in parsimony and model complexity that having a hybrid system does.

639 **6 Conclusion**

640 The present paper leaves open a number of puzzles – empirical and theoretical – that may be
641 fruitfully taken up in future work.

642 On the empirical side, although we have identified strong effects of N2 and compound fre-
643 quency in governing the rate of nasalization, it is not clear whether the claim by Ito & Mester that

644 the morphological status specifically (free vs. bound) itself matters, above and beyond the effect
645 of morpheme frequency. Answering this question would require a more targeted comparison of
646 N2s which are obligatorily bound with frequency-matched N2s which can be free-standing.

647 On the theoretical side, we have presented data that we think are important contributions to
648 the enterprise of phonological theory-building and comparison, and suggested that our results
649 favor those analyses that allow for individual morphemes to exhibit frequency-sensitive behav-
650 ior — whether accomplished representationally or computationally. However, we did not carry
651 out in-depth qualitative, let alone quantitative, comparison of the different classes of analysis
652 discussed in section 5; in the development of robust and psycholinguistically-informed phono-
653 logical theory, this is an equally important contribution that remains for future work. We hope
654 this paper serves as fuel for those who will carry it out.

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