

Modeling frequency-conditioned paradigm uniformity in Japanese voiced velar nasalization*

Canaan Breiss, University of Southern California[†]

Hironori Katsuda, University of Kansas

Shigeto Kawahara, Keio University/International Christian University

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Abstract

Recent quantitative work on the variable [g] ~[ŋ] alternation in compounds of certain dialects of Japanese has revealed *token frequency* of the compound as a whole, and of the compound's second-member (N2) in its freestanding form, to be important predictors of the alternation (Breiss et al., 2021a, Breiss et al. *to appear*). In this paper, we propose a formal phonological analysis of data presented in Breiss et al. (*to appear*) that integrates usage-based factors like frequency with the action of the phonological grammar, extending the mechanisms of lexicon-grammar interaction proposed in Breiss (2024). We demonstrate that the model fits experimental data better than—or at least comparably to—a theoretically-naïve statistical model proposed in the previous work. Based on the success of our modeling, we discuss the role of token frequency in phonological patterning more broadly, and how the mechanism that we propose in this paper might be extended to unify a range of contradictory frequency-dependent processes that have been observed in the literature.

1 Introduction

This paper is about how to integrate information about *usage frequency*—here, the token frequency of morphemes in the language experience of an individual speaker—into a constraint-

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[†]Contact: cbreiss@usc.edu

18 based phonological grammar formalism that characterizes that speaker’s generative linguistic
19 knowledge.

20 We take as our empirical case the frequency-conditioned variability in optional paradigm uni-
21 formity in voiced velar nasalization (henceforth “nasalization”) in phonologically-conservative
22 Japanese dialects, recently studied using corpus data by Breiss et al. (2021b) and experimentally
23 verified in Breiss, Katsuda and Kawahara (*to appear*) – henceforth BKK. These studies are the
24 latest in a long research tradition centered on the allophonic distribution of /g/ in conserva-
25 tive Japanese dialects, where a static phonotactic restriction enforces /g/ to be realized as [g]
26 prosodic-word-initially and [ŋ] elsewhere (e.g. Kindaichi 1942; Trubetsky 1969; Labrune 2012).
27 This correspondence is disrupted in compounds with /g/-initial second member (N2) that can oc-
28 cur as a free morpheme: in compounds with N2s that do not occur as free-standing words, the
29 /g/ → [ŋ] alternation is exceptionless, but in compounds where N2 may additionally occur as a
30 free-standing word (that is, with initial [g]) the nasalization process is optional (Ito and Mester,
31 1996, 2003).

32 The contribution of recent work by BKK (reviewed in detail in section 2) is to characterize this
33 variation in quantitative detail, and in particular to highlight how the token frequency of both
34 the compound and the free N2 impact the outcome of optional nasalization: higher frequency
35 compounds encourage more nasalization of medial /g/ to [ŋ], while higher frequency free N2s
36 encourage more retention of medial /g/ as [g], remaining uniform across the paradigm of their
37 free-standing forms and compound forms (Steriade, 2000; Benua, 2000).

38 The novel contribution of this paper is to provide a formally explicit model of the experimental
39 data. The model builds upon the Voting Bases model of lexicon-grammar interaction (Breiss,
40 2024), originally proposed to model Lexical Conservatism (Steriade, 1997). Lexical Conservatism
41 is a type of paradigm uniformity where the distribution of stem allomorphs (referred to as “bases”)
42 in a paradigm influences the way that paradigm accommodates new members. The canonical
43 example comes from Steriade (1997), who observed that the phonologically-similar forms *rémedy*
44 and *párody* differ in their behavior when affixed with *-able*, yielding *remédiable* with shifted stress,
45 but *párodiable*, with fixed stress. She argued that this difference stems not from the forms *remedy*
46 and *parody* themselves, but from the fact that *remedy* has a stem allomorph *remédi-* in *remédial*
47 that satisfies the marked lapse arising from affixation.

48 Breiss (2024) examined the same Lexical Conservatism dependency using novel derived forms
49 (like *lábor* + *-able*, with related form *labórious*, and *pláster* + *-able* with no phonologically-
50 advantageous related form), and found that in experimental settings, speakers are sensitive not
51 only to the *presence* of the phonologically-beneficial stem allomorph (like *remédial* and *labórious*),
52 but also to its salience in the lexicon as manipulated by priming. To account for these data, he
53 proposed a formal phonological model that integrates the influence of the contents of the lexi-

54 con along with their resting activation, enabling the phonological grammar to be sensitive to the
55 psycholinguistic properties of the morphemes which it manipulates. Breiss (2024) termed this
56 formal model of lexicon-grammar interaction the *Voting Bases* model.

57 In this paper, we demonstrate that the Voting Bases model extends, without modification, to
58 the separate case of lexicon-grammar interaction found in Japanese nasalization. The success
59 of the model suggests that the foundational principles of the Voting Bases model may be a good
60 candidate for a general theory of the way that the lexicon and grammar interact. This finding also
61 underscores the explanatory value to be gained for phonological phenomena by adopting a more
62 psycholinguistically-nuanced portrait of the lexicon as a dynamic substrate that can influence the
63 computations of the grammar on the items which it contains.

64 The layout of the paper is as follows: the first two sections of the paper review in some depth
65 basic facts about Japanese nasalization drawn from the literature (section 2), and then specifically
66 reviews in detail Experiment 1 of BKK (section 3). The following section, 4, focuses on the Voting
67 Bases model, and how we apply it to the context of optional paradigm uniformity. Section 5 then
68 actually fits the model to the experimental results, and discusses relative and absolute model fit in
69 comparison to minimally-different models that incorporate only some of the assumptions of the
70 Voting Bases model. The paper closes in section 6 with a discussion of broader issues, touching
71 on how such a system might come to be in the mind of the learner, on the merits of a joint model
72 of psycholinguistic and grammatical influence on word formation, and on what a unified theory
73 of token frequency effects on the phonological grammar might look like.

74 2 The traditional picture of Japanese nasalization

75 The data that we model in this paper comes from Experiment 1 of BKK, which investigated the
76 variation between [g] and [ŋ] induced by the phonotactics of phonologically-conservative di-
77 alects of Japanese. The pattern, which has been well-studied in both in generative and non-
78 generative literature on Japanese linguistics (Kindaichi, 1942; Trubetskoy, 1969; Hibiya, 1995;
79 Labrune, 2012; Ito and Mester, 1996, 2003), is exemplified in the complementary distribution of
80 [g] and [ŋ] shown in the monomorphemic data in example (1) below, where the voiced oral velar
81 stop is only permitted word-initially, and the velar nasal is only permitted word-medially.

- 82 (1) a. /kaŋami/ → [kaŋami]
83 “mirror”
84 b. /gimu/ → [gimu]
85 “obligation”

86 We assume throughout that non-alternating forms are stored surface-true as URs in the lexi-

87 con, in accordance with the phonological tradition of (Strong) Lexicon Optimization (Prince and
 88 Smolensky, 1993; Sanders, 2003). This stance is supported by psycholinguistic research on the
 89 contents of lexical representations, reviewed in section 4.1.

90 Japanese’s extensive use of compounding in word-formation gives the opportunity for the
 91 phonotactic restriction to drive alternations, seen in examples (2)-(5) below. Here we see that
 92 when a /g/-initial morpheme is word-initial (either as a prosodically-free word, in examples (2)–
 93 (4), or as the first member (N1) of a compound, in example (5)¹), it is realized with an initial [g],
 94 while when it occurs as the second member of a compound (N2) it is realized with initial [ŋ].
 95 Critically for the current study, Ito and Mester (1996) observed that although in all cases the /g/-
 96 initial N2 *may* be realized word-medially with initial [ŋ], nasalization is optional when the N2
 97 can stand on its own as a prosodically-free form (cf. the “b” series in (2)–(4) vs. (5c))—a case of
 98 optional paradigm uniformity.

- 99 (2) a. /hai + gan/ → [hai-ŋan] ~ [hai-gan]
 lung cancer
 100 “lung cancer”
 101 b. /gan/ → [gan]
 102 cancer
 103 “cancer”
- 104 (3) a. /noo + geka/ → [noo-ŋeka] ~ [noo-geka]
 brain surgery
 105 “brain surgery”
 106 b. /geka/ → [geka]
 107 surgery
 108 “surgery”
- 109 (4) a. /doku + ga/ → [doku-ŋa] ~ [doku-ga]
 poison moth
 110 “poison moth”
 111 b. /ga/ → [ga], “moth” (a free-standing morpheme)
- 112 (5) a. /doku + ga/ → [doku-ŋa], *[doku-ga]
 poison fang
 113 “poison fang”
 114 b. /ga + ʒoo/ → [ga-ʒoo]
 fang castle
 115 “main castle”

¹We temporarily adopt here for the traditional assumption that the [g]-initial form of a free N2 is underlying, for expository ease and continuity with the previous literature. Our own proposal is laid out in section 4.

116 c. /ga/ → *[ga] (a bound morpheme)
 fang
 117 “fang”

118 Breiss et al. (2021b) examined this variation in a corpus derived from a pronouncing dictionary
 119 (NHK, 1993) and found that among compounds with free N2s, the two most prominent predictors
 120 of whether an item would be nasalized was the frequency of the N2’s free [g]-initial form, and
 121 the frequency of the whole compound. These effects ran in opposite directions: higher frequency
 122 compounds were more likely to be nasalized (the left facet of Figure 1); on the other hand, the
 123 more frequent the free N2, the less likely the nasalization was (the right facet of Figure 1).

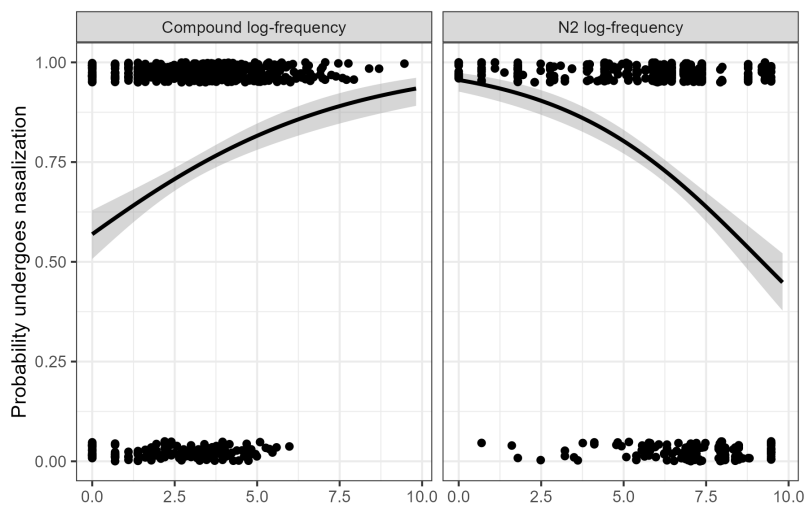


Figure 1: The effects of whole compound frequency (left facet) and N2 frequency (right facet) on the probability of nasalization (vertical axis), with binomial smooths in the corpus data. One dot represents one lexical item; vertical jitter has been added for readability. Figure and caption adapted with permission from Breiss et al. (*to appear*), data from Breiss et al. (2021b).

124 The corpus data was modeled as a case of probabilistic paradigm uniformity in Breiss et al.
 125 (2021a) using Output-Output Faithfulness constraints (Benua, 2000) indexed to items binned by
 126 the relative frequency of each compound and N2. The paper was limited, however, by the untested
 127 assumption of their model that the frequency-modulation of paradigm uniformity in their corpus
 128 data actually represents the synchronic knowledge of speakers. Additionally, their formal model
 129 was not explicitly informed by psycholinguistic considerations and thus its linking hypothesis
 130 between frequency (necessarily a lexical characteristic) and the phonological grammar was vul-
 131 nerable to criticism on grounds of being stipulative—in other words, there was nothing in their
 132 model that prevented the opposite relation between frequency and paradigm uniformity from
 133 holding.

134 In this paper, we offer two improvements on the state of affairs in Breiss et al. (2021a). First, we
135 model experimental data from Breiss et al. (*to appear*) (BKK) where the frequency-conditioning
136 of the variable paradigm uniformity is reproduced in existing compounds and extended to novel
137 ones. Second, we do this by extending the Voting Bases model of Breiss (2024) which is compat-
138 ible with consensus understanding of the way lexical frequency is connected to the lexical rep-
139 resentation and activation, and which offers an explicit linking hypothesis relating the real-time
140 dynamics of the lexicon to the representation and computations of the phonological grammar.

141 **3 BKK’s Experiment 1**

142 BKK carried out two experiments on Japanese nasalization. Their paper had the goal of seeing
143 whether the corpus data was representative of speakers’ generalizable knowledge, both in the
144 aggregate and also at the level of the individual. They found that both individually and in aggre-
145 gate, speakers’ propensity to nasalize displayed sensitivity to the frequency of the free N2 and
146 compound, in existing and novel compounds. In this paper, we focus our modeling efforts on the
147 results of their Experiment 1, which we describe in some detail below.²

148 **3.1 Stimulus selection**

149 BKK chose stimuli that were roughly balanced between existing Japanese compounds of varying
150 frequencies, and novel (that is, zero-frequency) semantically-compositional compounds. Both
151 existing and novel stimuli had attested free N2s of a range of frequencies. Out of a desire to
152 sample compounds with a wide range of frequencies that would likely be known to participants,
153 existing compounds ranged from two to eight moras in length, while all novel compounds were
154 four moras long.

155 **3.2 Participants**

156 BKK sought to recruit speakers of the phonologically-conservative Tōhoku dialect of Japanese,
157 and used word-of-mouth and snowball sampling to find 20 speakers. In order to increase the
158 precision of individual-level estimates of experimental manipulations, all but one speaker par-
159 ticipated in the experiment twice, with the two sessions separated by a period ranging from a
160 few weeks to a few months. Although the phonologically-conservative Yamanote dialect was

²They also sought to determine whether correlation between nasalization and the overall prosodic size of the compound, which is observed in the corpus (Breiss et al., 2021b) but is a typologically unusual pattern, was replicated in participants’ online productions (Experiment 2). They actually found that there was no evidence of a direct relationship between nasalization and global prosodic length (cf. Jiang 2023). We therefore do not address this experimental data here, as our point is made in the simpler case of data from Experiment 1.

161 the subject of the corpus study of Breiss et al. (2021b) and of much of the previous linguistic
162 analysis, this dialect is currently spoken only by the very elderly who might not be comfortable
163 participating in a study online.

164 Before participating in the experiment, each participant was screened with a short dialect
165 questionnaire to ensure that their speech exhibited the allophonic distribution of word-initial
166 [g] and word-medial [ŋ]. This was done because the literature addressing nasalization assumes
167 that it is triggered in order to enforce compliance with this phonotactic restriction imposed on
168 monomorphemic words; therefore it is important to base conclusions about the variability of
169 nasalization on data from speakers who do exhibit the phonotactic in question.³ For the purposes
170 of the model which we develop, we will see that these monomorphemic words provide crucial
171 evidence for the lower bound of the weight of the markedness constraint driving nasalization,
172 since with data from compounds alone, it is not uniquely identified against the background of
173 faithfulness constraints that the Voting Bases model uses (see section 5 for further details).

174 The dialect questionnaire consisted of a production task where speakers were asked to read
175 aloud 10 monomorphemic words with word-initial [g] of varying frequencies, and 10 monomor-
176 phemic words with word-medial [ŋ]. The stimuli were written with *kanji* orthography, which
177 does not distinguish between [g] and [ŋ]—this is also true of the main production experiment
178 described below, so we follow BKK in assuming that the participants’ production was not influ-
179 enced by orthographic factors. The twenty words were shown to the participant in a random
180 order, and their productions were recorded; only the eight participants who exhibited the target
181 pattern of allophony in all monomorphemes were invited to participate in the main experiment.

182 3.3 Experiment structure

183 Each experimental session proceeded as follows. First, participants were asked to complete the
184 dialect questionnaire; if they met the criterion discussed above, they were invited to complete
185 the rest of the experiment. Before the production task, participants read out loud and indicated
186 whether they knew half of the free N2s in the experiment—this was done in order to prime them,
187 under the hypothesis that raising their resting frequency in the participant’s lexicon would in-
188 fluence their phonological behavior (see further discussion in section section 4.4). After this
189 knowledge check, participants saw each compound one at a time in a random order, and pro-
190 duced the form aloud while their speech was recorded. After the production task, participants
191 produced and indicated knowledge of the other half of the free N2s in the experiment, as well
192 as all of the compounds. The N2s that were primed were counter-balanced between participants
193 across the two runs of the experiment.

³The question of how nasalization is produced or represented by speakers who entirely lack, or only variably enforce, the [g] ~[ŋ] allophony in monomorphemic words is not addressed by BKK, nor do we consider it here.

194 **3.4 Results**

195 BKK found that the participants reflected at an individual level the frequency-conditioned vari-
196 ability seen in the corpus study of Breiss et al. (2021b). In existing compounds (Figure 2), their
197 productions were influenced by both the frequency of the compound (the left facet), for which
198 higher values correlated with more nasalization, and by the frequency of the free N2 (the right
199 facet), where higher values correlated with less nasalization.

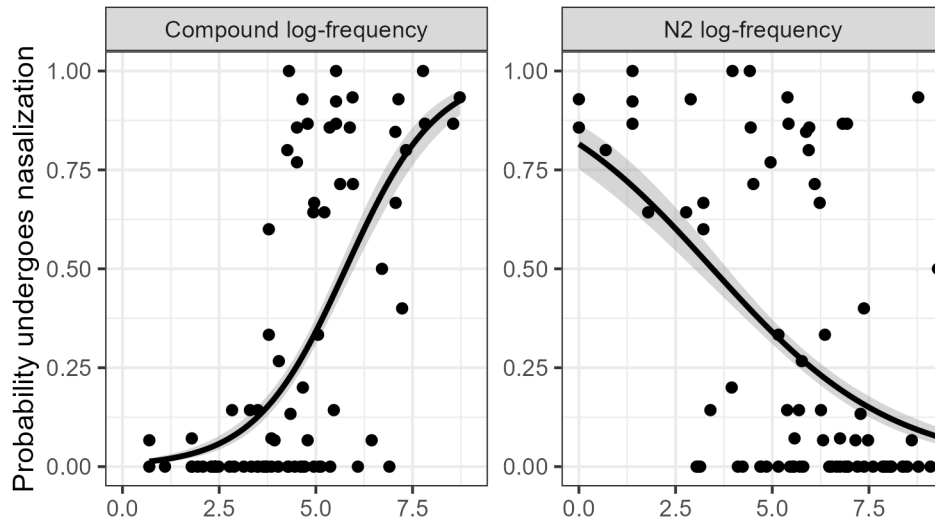


Figure 2: Probability of nasalization (the vertical axis) plotted against compound log-frequency (the left facet) and N2 log-frequency (the right facet), with binomial smooths for readability, in the experiment by BKK. Plot and caption reproduced with permission from Breiss et al. (*to appear*).

200 Figure 3 plots the same effect of N2 frequency in novel compounds: forms with higher-
201 frequency N2s were less likely to undergo nasalization relative to those with lower-frequency
202 N2s.

203 Finally, BKK found that the frequency effect was stable at the level of the individual, across
204 existing and novel compounds, which is plotted in Figure 4. In this Figure, the horizontal axis plots
205 the strength and direction of the effect of N2 log-frequency in novel compounds, and the vertical
206 axis plots the strength and direction of the effect of N2 log-frequency in the existing compounds;
207 see the caption of Figure 4 for further details. Although different participants were more or less
208 sensitive to the frequency of a given N2, lying higher or lower on each axis, there was uniformity
209 in this degree of sensitivity such that the two co-varied along a diagonal line through the center
210 of the plot. BKK interpreted this correlation as evidence that morpheme usage frequency and
211 phonological markedness have separable, distinct influences on speaker productions.

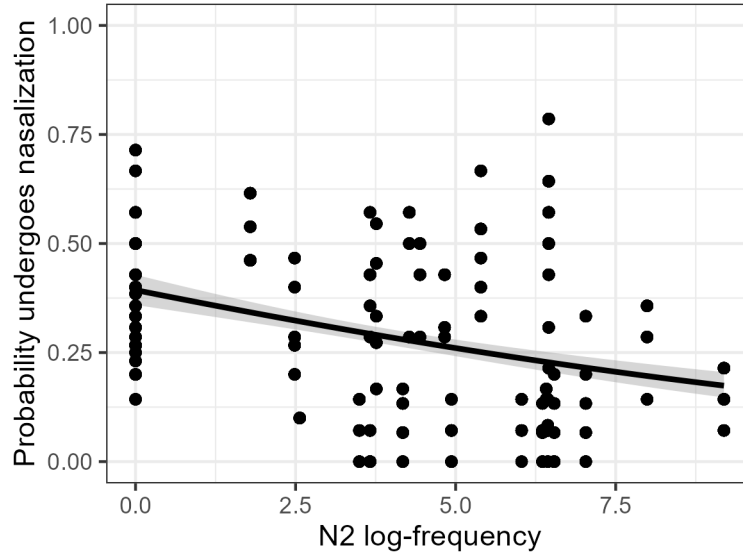


Figure 3: The probability of undergoing nasalization in novel compounds, plotted against N2 log-frequency, with a binomial smooth to aid readability. Plot and caption reproduced with permission from Breiss et al. (*to appear*).

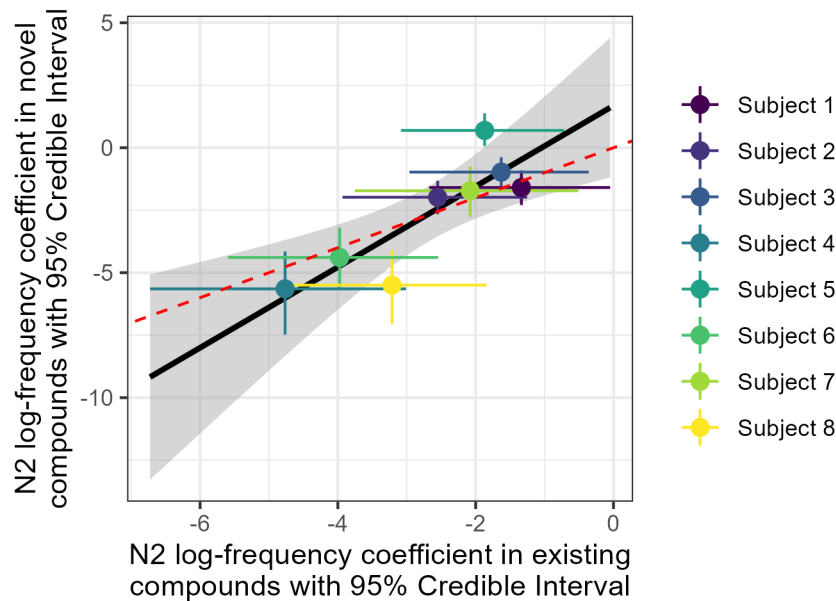


Figure 4: The coefficient of N2 log-frequency in novel compounds, derived from the model in Table 1 of Breiss et al. (*to appear*), is plotted on the horizontal axis, and the coefficient for N2 log-frequency in existing compounds, derived from the model summarised in Table 3 of Breiss et al. (*to appear*), is plotted on the vertical axis. Points represent median values of the posterior with ranges encompassing the 95% Bayesian Credible Intervals, colors represent speakers, and a linear smooth has been added for readability, with the line of slope 1 intersecting the origin in dotted red. Plot and caption adapted with permission from Breiss et al. (*to appear*).

212 **3.5 Summary and goals for modeling**

213 To summarize, the findings of BKK that are relevant for the modeling task of this paper are
214 the following. Among those speakers for whom the phonotactic restriction enforcing [g]~[ŋ]
215 allophony was exceptionless in monomorphemic words:

- 216 1. Phonotactically-driven nasalization is variable in compounds with prosodically-free N2s.
- 217 2. In these compounds, the probability of nasalization is increased by higher compound fre-
218 quency, and decreased by higher N2 frequency.
- 219 3. The frequency effect is uniform within individuals across existing and novel compounds.

220 Below, we propose a formal model of these facts, using the Voting Bases model to relate a lexicon
221 containing usage-frequency information to a phonological grammar couched in the Maximum
222 Entropy (MaxEnt) framework.⁴

223 **4 Modeling token frequency in the phonological grammar**

224 Based on the facts laid out above, we seek a model of the phonological grammar that allows
225 non-phonological properties of individual lexical items (here, frequency) to influence their par-
226 ticipation in phonological processes (here, paradigm uniformity). Note that we specifically aim
227 to model phonological and non-phonological influences on the outputs of the phonological gram-
228 mar, rather than any possible morphological or paradigmatic effects on phonetic realization (see
229 Purse et al. (2022) for a review), about which the Voting Bases model as laid out in Breiss (2024)
230 makes no predictions.

231 **4.1 The contents of a lexical entry**

232 As prolegomena to the grammatical model, it will be important to establish some relevant context
233 regarding the contents of the lexicon, because it is these representations that are at stake in
234 discussions of token frequency. Psycholinguistic research has amassed a large body of evidence
235 that the lexicon is richly structured, with numerous types of linked representations of various
236 levels of detail grouped under the same lexical entry. We do not review this research in depth
237 here, but simply highlight the findings relevant to developing the type of integrated phonological

⁴We do not attempt to model the frequency of the first compound member, N1, on the probability of nasalization in compounds, since this was not manipulated by BKK. Future work might profitably pursue this question experimentally and formally, since corpus data in Breiss et al. (2021b) suggests that higher N1 frequency may also independently lower the probability of nasalization; see Rebrus and Törkenczy (2017) for a similar finding of N1 frequency on compound coherence in Hungarian vowel harmony.

238 theory referenced above. For a thorough discussion and literature review on the (phonologically-
239 relevant) contents of a lexical entry, see Pierrehumbert (2016); for more on how this information
240 interacts with the Voting Bases theory in cases beyond those relevant for the nasalization, see
241 Breiss (2021, 2024).

242 Since nasalization concerns paradigm uniformity, we assume the lexical entry for an exist-
243 ing word lists (among many other things) their allomorphs (cf. Strong Lexicon Optimization,
244 Sanders 2006): for a non-alternating monomorpheme like [kaŋami] “mirror”, this would be sim-
245 ply /kaŋami/; for a monomorpheme that can appear as an N2 and undergo nasalization, such as
246 [ga]–[ŋa] “moth”, the lexical entry would list both /ga/ and /ŋa/. Finally, we assume that existing
247 compounds are stored whole, with nasalization applied so as to respect the phonotactic in the
248 lexicon (Albright, 2008; Martin, 2007).

249 With regard to non-phonological characteristics of the lexicon, we follow a large body of
250 evidence that lexical representations have differing degrees of salience or strength of encoding,
251 which is often referred to as their *resting activation* (Morton, 1970). Following Breiss (2021, 2024),
252 we take resting activation to correspond to the strength of a memory representation itself, not a
253 number or rank stored in long-term memory as a characteristic of the lexical item. Thus, char-
254 acteristics (long-term or dynamic) of lexical items like their frequency, and whether or not they
255 were recently activated (for example, by priming), all contribute dynamically to an item’s resting
256 activation. Importantly, also following Breiss, we use the term “resting activation” as a stand-
257 in for any scalar summary statistic that can be derived from an implemented model of lexical
258 dynamics. We remain intentionally agnostic as to the specific model of these dynamics, simply
259 stressing that so long as such a model can be used to drive a measure of relative salience influ-
260 enced by the factors just mentioned, the Voting Bases model can make reference to it to scale
261 faithfulness constraint violations. We discuss how resting activation is modeled as influencing
262 the phonological grammar below in section 4.4.

263 4.2 The Voting Bases model

264 We now turn to a formal phonological model of the Japanese nasalization data. We use the Voting
265 model of Base competition proposed in Breiss (2021, 2024). The Voting model has been used to
266 model data in Lexical Conservatism in English and Spanish, and is broadly compatible with the
267 view of the lexicon laid out above. Here, we extend the model to the probabilistic paradigm-
268 uniformity found in Japanese nasalization.

269 The Voting Bases model has two parts: the first is that all listed stem allomorphs (“bases”)
270 in the lexicon exert an analogical pull on derivatives (operationalized using allomorph-specific
271 faithfulness constraints), violations of which are scaled in proportion to the resting activation of
272 the representation to which faithfulness is being assessed. The second part is that markedness

273 constraints evaluate candidates in the standard way for constraint-based phonological models.

274 The Voting Bases model assumes a probabilistic, weighted-constraint phonological grammar;
275 here, we use MaxEnt Harmonic Grammar (Smolensky, 1986; Goldwater and Johnson, 2003), but
276 in principle we could also use another grammar formalism that has these characteristics, like
277 Stochastic (or Noisy) Harmonic Grammar (Boersma and Pater, 2016). We use MaxEnt since it has
278 various strengths; e.g. it directly relates Harmony to probability (Hayes, 2022), permits constraint
279 cumulativity by default (Jäger and Rosenbach, 2006; Breiss, 2020), has a learning algorithm to set
280 its weights, and is rooted in well-understood statistical techniques used widely outside linguistics
281 (Jurafsky and Martin, 2009, ch. 5). We stress, however, that our analyses can be recast in terms
282 of other stochastic constraint-based frameworks.

283 4.3 Constraints

284 In the analysis developed in this paper, we adopt the general approach of Ito and Mester (1996,
285 2003), following loosely Breiss et al. (2021a). We only use three constraints: a single marked-
286 ness constraint to motivate nasalization (extending the spirit of the constraint *VgV from Ito and
287 Mester (2003) to be compatible with nasal-final N1s, which pattern identically to vowel-final N1s),
288 and a pair of faithfulness constraints which correspond to the second member’s free form and to
289 the analogical pull of the compound as a whole, if one exists. They are listed below.⁵

- 290 • ***INTERNAL-[g]**: Assign one violation for each word-internal [g] in a candidate.
- 291 • **ID-[nasal]-N2**: Assign one violation for each segment in the listed allomorph for the free-
292 standing N2 that does not match its corresponding segment in the feature [nasal].
- 293 • **ID-[nasal]-COMPOUND**: Assign one violation for each segment in the listed allomorph for
294 the full compound that does not match its corresponding segment in the candidate in the
295 feature [nasal].

296 Note that the constraint definitions do not make reference to scaling or the contents of the
297 lexicon; the proposal in the Voting Bases model is an architectural proposal about how psycholin-
298 guistic, “extra-grammatical” factors act within and beside the phonological grammar to influence
299 certain variable phenomena.

⁵The first faithfulness constraint plays the same role as faithfulness to the Remote Base in an analysis of Lexical Conservatism. The second faithfulness constraint parallels faithfulness to the Local Base in a Lexical Conservatism analysis (Breiss, 2021, 2024). We use more transparent names here for the sake of clarity, since nothing in the Voting Bases model structurally prioritizes Remote Bases over Local ones.

300 4.4 Modeling resting activation

301 The discussion in 4.1 above left open how a specific numerical value for resting activation might
302 be calculated on the basis of the psycholinguistic characteristics of item’s lexical entry. Here,
303 we take the approach of modeling it as function of log-frequency of the frequency of the allo-
304 morph, which is passed through the sigmoid function $\frac{1}{1+e^{-\log freq}}$ that translates the linear predic-
305 tor (i.e. $-\log freq$) into the bounded interval of $\{0,1\}$, which will be the scaling factor applied to
306 faithfulness violations. This is illustrated in Figure 5. The effect of this non-linear transforma-
307 tion will be to preserve the idea that it is less penalized to be unfaithful to low-frequency lexical
308 items compared to higher-frequency ones, while damping down the difference between extreme
309 values of the scale and rendering it bounded. The final move we make here is rather than using
310 *raw* log-frequencies, we use *scaled and centered* log-frequencies, following the statistical analysis
311 in BKK. This corresponds to the notion that it is not so much the *absolute* frequency of each item
312 that is important, but how frequent it is relative to the other competitor items in the lexicon (here
313 approximated by the population of items in the experiment), which is in line with previous work
314 on morphological decomposition in stored forms (Hay, 2001). Finally, in the analysis that we
315 develop below, we do not model the priming of N2, since BKK did not find substantial evidence
316 that it affected their experimental data.⁶

317 4.5 Schematic illustrations

318 Before modeling the experimental data itself, it will be useful to work with some toy data to get
319 a feel for how resting-activation-scaled faithfulness violations interact with the dynamics of a
320 MaxEnt grammar. First, let us consider the case of novel compounds, since they are the simplest
321 case to lay out the workings of the analysis. Recall the empirical pattern: here, although the
322 frequency of the compound is zero, we nevertheless find that nasalization is modulated by the
323 frequency of N2. Now, consider the case of two hypothetical novel compounds, one with a higher-
324 frequency N2, and one with a lower-frequency N2, such that when the sigmoid transformation
325 is applied to their frequencies the higher-frequency form scales its violations of ID-[nasal]-N2 by
326 0.7, and the lower scales its own violations of the same constraint by 0.3 (these specific numbers
327 are chosen purely for the sake of illustration). Using the constraints defined in section 4.3 above,
328 we can define the tableaux below in Figure 6.

⁶The Voting Bases framework is easily extensible to multiple predictors of resting activation: to incorporate priming, one could simply treat the term passed into the sigmoid as itself a log-linear model, adding a coefficient (weight) for the effect of priming, in addition to a coefficient for the effect of lexical frequency. This is beyond the current scope of this paper, however, and so we simply assume a fixed coefficient for lexical frequency, since there being only one predictor in the log-linear model for resting activation would make the coefficient of frequency redundant with the weight of the faithfulness constraint being scaled. Similarly extensions of the Voting Bases model could also model by-participant variability in the priming effect using a hierarchical model structure.

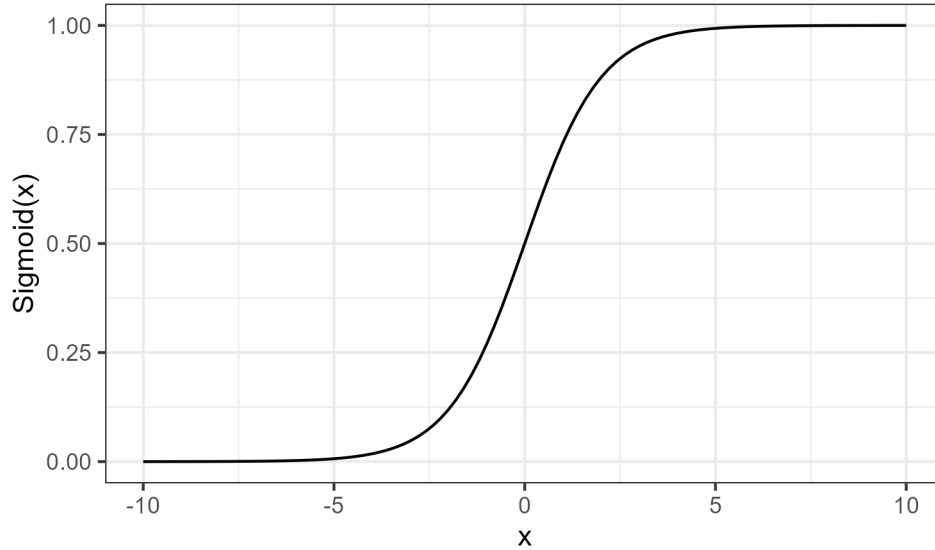


Figure 5: Sigmoid function that translates the (centered) frequencies into the scaling factors. See text for details.

$/\dots/N1, /g\dots/High-freq.N2$ Weight:	*INTERNAL-[g]	ID-[nas] _{N2}	H	<i>p</i>
a. [...g ...]	2	1	2	.21
b. [...ŋ ...]	1	.7	.7	.79
$/\dots/N1, /g\dots/Low-freq.N2$ Weight:	*INTERNAL-[g]	ID-[nas] _{N2}	H	<i>p</i>
c. [...g ...]	2	1	2	.15
d. [...ŋ ...]	1	.3	.3	.85

Figure 6: Schematic application of the Voting model of Base Competition to the formation of a novel compound in the *wug*-test.

329 We can see that the pull of faithfulness to the N2 with higher frequency is stronger than
 330 the one with lower frequency, though both are relatively marginal outcomes since the weight of
 331 *INTERNAL-[g] dominates the distribution of probabilities in this scenario.

332 Moving on to existing compounds, we now must add another item to the lexical entry we are
 333 considering in our left-hand input cell to our tableaux, shown in Figure 7. For the sake of minimal
 334 contrasts, we assume that the frequency of both N2s are equal and medial relative to the examples
 335 in Figure 6 above, allowing us to examine the effect of compound frequency holding N2 frequency
 336 constant. However, in our analysis of the actual data, both scaling factors are independently set
 337 on a per-item basis.

338 Here we see that the scaling of the compound again depends on frequency, but because of the

$/\dots/N_1,$ $/\dots\eta\dots/_{High-freq.compound}$	$/g\dots/N_2,$ Weight:	*INTERNAL-[g]	ID-[nas] $_{N_2}$	ID-[nas] $_{Compound}$	H	p
e. [...g ...]	1			.7	2.7	.09
f. [...η ...]			.5		.5	.91
$/\dots/N_1,$ $/\dots\eta\dots/_{Low-freq.compound}$	$/g\dots/N_2,$ Weight:	*INTERNAL-[g]	ID-[nas] $_{N_2}$	ID-[nas] $_{Compound}$	H	p
g. [...g ...]	1			.3	2.3	.14
h. [...η ...]			.5		.5	.86

Figure 7: Schematic application of the Voting model of Base Competition to the formation of an existing compound in the *wug*-test.

339 assumption we made about the listed form of the compound—specifically, that phonologically
340 well-formed words are preferentially the target of lexicalization (Albright, 2008; Martin, 2007)—
341 we find that the faithfulness to the compound’s UR penalizes the candidate that does not exhibit
342 nasalization and violates markedness.

343 Finally, we lay out the case where the competition between candidates is driven primarily by
344 faithfulness. Above, where markedness had a high weight, the candidate that satisfied marked-
345 ness had a higher probability than the one which violated it, and the effects of the faithfulness
346 constraints were on the probability of the minority candidate. In the scenario where marked-
347 ness is low and the weights of the faithfulness constraints are dominant, the majority candidate
348 is the one that satisfies faithfulness to the whole compound, and the presence of the N2 is the
349 main reason that the unfaithful (but markedness-satisfying) candidate gets appreciable probabili-
350 ty; this is a type of “analogical” effect where markedness has little role, as in Figure 8, in which
351 the markedness constraint is assigned a very low weight (here, arbitrarily set as 0.1).

352 5 The model in action

353 Moving on to the analysis itself, we first fit models separately to the existing and novel com-
354 pound data, to demonstrate the suitability of the Voting model in each context, and also to allow
355 better comparison to the statistical models fit to the experimental data above. We then consider
356 how nasalization might be modeled more comprehensively by incorporating information from
357 non-alternating monomorphemes where the complementary distribution between [g] and [ŋ]
358 is enforced, and also by incorporating our knowledge that the free form of N2s surface non-
359 alternatingly with initial [g], despite the presence of an [ŋ]-initial stem allomorph.

360 In all cases, we fit the MaxEnt models using the *Solver()* function in Microsoft Excel (Fylstra

$/\dots/N_1,$ $/\dots\eta\dots/_{High-freq.compound}$	$/g\dots/N_2,$ Weight:	*INTERNAL-[g]	ID-[nas] $_{N_2}$	ID-[nas] $_{Compound}$	H	p
i. [...g ...]	1			.7	1.5	.27
j. [...η ...]			.5		0.5	.73
$/\dots/N_1,$ $/\dots\eta\dots/_{Low-freq.compound}$	$/g\dots/N_2,$ Weight:	*INTERNAL-[g]	ID-[nas] $_{N_2}$	ID-[nas] $_{Compound}$	H	p
k. [...g ...]	1			.3	0.7	.45
l. [...η ...]			.5		0.5	.55

Figure 8: Schematic application of the Voting model of Base Competition to the formation of an existing compound in the *wug*-test, in a regime where faithfulness is strong and markedness weak.

et al., 1998), and used a relatively weak Gaussian prior of Normal(0,10) on constraint weights, which has the effect of allowing weights to vary in response to values that best fit the data, while making extreme values (here, above twenty or so) less appealing. For more on priors on weights in MaxEnt phonological models, see Wilson (2006) and White (2017). All models fit in this paper are provided in the supplementary materials.

5.1 Existing compounds

We first applied the analysis sketched in section 4.5 to data from existing compounds. Recall that in these forms, compounds with higher-frequency N2s are more likely to resist nasalization than those with lower-frequency N2s, but that compound frequency itself also influences nasalization, with higher-frequency compounds favoring the surface-realization of their underlying [η]. We take as our data to model the counts of compounds produced having undergone nasalization or not, in cases where speakers know both the compound and the N2 in question.

The best-fitting constraint weights for existing compounds are 6.93 for ID-[nasal]-COMPOUND, and 7.22 for ID-[nasal]-N2, with *INTERNAL-[g] receiving a weight of zero. The weights of the two faithfulness constraints were not significantly different from one another, as assessed via a likelihood ratio test: $\Delta\log\text{-likelihood} = 1.3$, $p = 0.11$; a similar conclusion was suggested by the near-zero difference in the sample-size corrected AIC of the two models: $\Delta\text{AICc} = 1.8$. AICc differences greater than 10 are typically taken to indicate strong support for the model with the lower AICc value; for more on model-comparison in statistical models and phonological grammars, see Shih (2017) and Wilson and Obdeyn (2009). This result suggests that the attractive influence of both Bases is critical in driving the alternation in attested forms; the zero weight of the markedness constraint *INTERNAL-[g] indicates that in existing compounds, analogical faith-

383 fulness is doing all the work here, despite the assumption in the literature that the alternation is
384 markedness-driven. We will revisit the role of markedness when evaluating the joint model of
385 novel and existing compounds below in section 5.3.

386 We also compared the full model to one where the two faithfulness constraints were allowed
387 to take on different values but were not scaled by frequency. As one might expect, since low- and
388 high-frequency forms have the same violation profiles in the phonological grammar, a grammar
389 without access to frequency information can only predict one rate of nasalization across all forms;
390 this model fits the data dramatically less well ($\Delta\log\text{-likelihood} = 264.87$, $p < .001$ with one degree
391 of freedom, $\Delta\text{AICc} = 527.15$).

392 Finally, we evaluate the absolute performance of the model by examining how well it fits
393 the data: although the two models have different internal structures, we can ask whether the
394 theoretically-informed MaxEnt model here does as good a job in explaining the data patterns
395 as the theory-neutral mixed effects logistic regression model reported by BKK when assessing
396 the statistical robustness of the experimental results.⁷ We do this using the measure of R^2 , which
397 ranges from zero to one, and can be thought of as the proportion of the variation in the dependent
398 variable (here, whether nasalization applies or not) explained by the collection of independent
399 variables (the phonological and lexical characteristics of interest).

400 We used the *r2_bayes()* function from the *performance* package (Lüdtke et al., 2021) to obtain
401 the marginal R^2 of the statistical model—that is, the amount of variance in the data explained by
402 the fixed effects—and compared it to the R^2 for the MaxEnt model.⁸ Since the statistical model
403 is Bayesian, we obtain a median and 95% Credible Interval for our R^2 : 0.48 and [0.31, 0.56], re-
404 spectively. This is lower, though still relatively comparable, to the MaxEnt models R^2 of 0.63, for
405 which we have only a point estimate. Although the two are relatively close, the point value for
406 the marginal R^2 of the MaxEnt model is outside the 95% Credible Interval of the statistical model;
407 this comparison suggests that the theoretically-structured model out-performs the theory-blind
408 statistical one. While we find this result to be encouraging, this conclusion is tentative, however—
409 since the MaxEnt model does not capture variation at the level of the speaker, it may be overfitting
410 the data somewhat, attributing to the population grammar variance that should more conserva-
411 tively be attributed to speaker-level idiosyncrasies.

⁷The model specification in BKK was as follows: $\text{Nasalization} \sim 1 + \text{LogN2Freq} * \text{N2Primed} + \text{LogCompoundFreq} + \text{NasalFinalN1} + (1 + \text{LogN2Freq} * \text{N2Primed} + \text{LogCompoundFreq} + \text{NasalFinalN1} \mid \text{Speaker}) + (1 + \text{N2Primed} \mid \text{Compound})$; see BKK section 3.3 and 3.4.1 for details.

⁸We used marginal R^2 , which makes reference to fixed effects only, since the conditional R^2 that takes into account the variance explained by both fixed and random effects has no direct comparison in the MaxEnt model we fit. For more on the relationship between mixed effects models and hierarchical structures in linguistic data, see Zymet (2019).

5.2 Novel compounds

Turning to entirely novel compounds, we fit the analysis sketched in section 4.5 to the data obtained by BKK. Here, the best fit weight for the constraints gives only ID-[nasal]-N2 a nonzero weight, at 2.18; faithfulness to existing compounds receives no weight (as expected, since novel compounds by definition do not have an existing compound to be faithful to), and also the markedness constraint is weighted zero. This is also to be expected, since the one degree of freedom in the model is actually the harmony *difference* between faithful and repairing compounds, and so since faithfulness receives a non-zero positive weight, the Gaussian prior prefers the other constraint to remain at zero.

As above, we compared the fit of the MaxEnt model of novel compounds to one that did not allow access to frequency information, which fit the data much less well relative to a model that does allow the lexicon to scale violations of faithfulness constraints ($\Delta\log\text{-likelihood} = 11.12$, $p < .001$ with one degree of freedom, $\Delta\text{AICc} = 22.19$). We also compared the absolute fit of our theoretically-motivated MaxEnt model to the purely statistical model fit by BKK⁹, and find that the R^2 of our model, 0.11, falls within the 95% Credible Interval of the median of that of the statistical model, 0.06 [0.00, 0.18]. While low in absolute terms, it is reassuring that it is in line with the statistical model, suggesting that our grammatical model is not overfit to the data. We suspect that the cause of the poor model fit may be that there is greater between-individual variation in novel compounds than in existing ones, while the model we fit to the novel data has fewer parameters, and thus is less expressive, than that fit to existing data.

5.3 Fitting a joint model

To get a more holistic picture of nasalization, we fit a joint model to both existing and novel compounds, and integrated the fact that the participants were included in the experiment on the basis of exhibiting complementary distribution of [g] and [ŋ] in monomorphemes. Therefore, in addition to both sets of compound data, the model included the monomorphemes used in the dialect questionnaire to screen participants for inclusion in the experiment, including frequency-based scaling of their faithfulness violations. Since we assume lexicon optimization, we cannot accurately assess the weight of the markedness constraint *INTERNAL-[g] that drives the language-wide phonotactic because the number of monomorphemes that we surveyed was relatively small in comparison to all the data in the Japanese lexicon that exhibits this complementary distribution (Ito and Mester, 2003), and thus contributes to the actual weight of *INTERNAL-[g] in speakers' grammars. However, we can find a lower bound on its weight by constraining the sets of weights

⁹Model specification: $\text{Nasalization} \sim 1 + \text{LogN2Freq} * \text{N2Primed} + (1 + \text{LogN2Freq} * \text{N2Primed} \mid \text{Speaker}) + (1 + \text{N2Primed} \mid \text{Compound})$; see BKK section 3.3 and 3.4.2 for details.

444 we consider to those that maximize the likelihood of the compound data, while simultaneously
 445 preserving allophony in monomorphemes (operationalized as having 95% or greater probability
 446 of faithful realization). The final model yielded weights listed in Table 9, and predictions plotted
 447 in Figure 10.

Constraint	Weight
*INTERNAL-[g]	0.16
ID-[nasal]-COMPOUND	5.12
ID-[nasal]-N2	4.25

Figure 9: Best-fitting weights for the experimental data, existing and novel compounds combined, that preserves the allophony in monomorphemes.

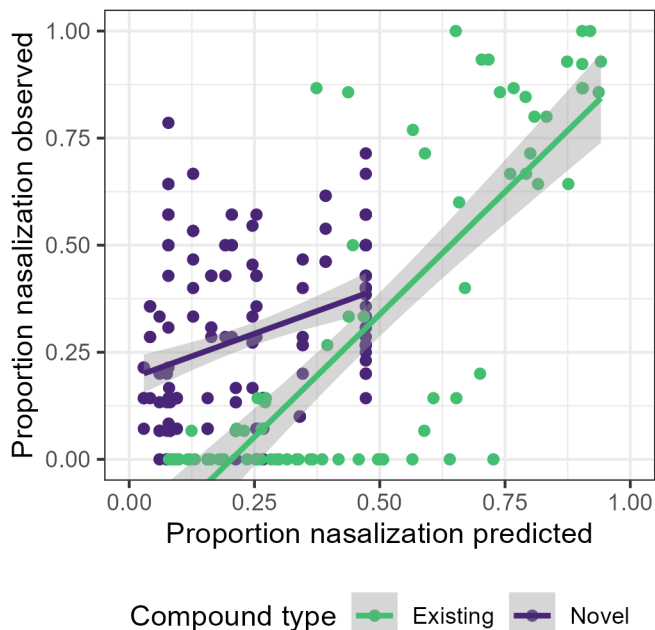


Figure 10: Predicted (vertical axis) vs. observed (horizontal axis) rates of nasalization for categories existing (green) and novel (purple) compounds under the combined model (weights in Table 9).

448 Although the integration of non-alternating monomorphemic words was intended to give a
 449 better picture of the weight of the markedness constraint, their inclusion ended up leaving the
 450 weights of the IDENT-[nasal] constraints relatively unchanged, with the markedness constraint
 451 only getting a modest weight. We take this to indicate that much of the variability seen in the
 452 experimental data is actually driven by a strong analogical effect of faithfulness, rather than
 453 paradigm-uniformity being parasitic on markedness. This is consistent with the weights ob-
 454 tained when fitting the individual datasets above; when forced to account for the non-alternation
 455 of monomorphemes, a modest weight for *INTERNAL-[g] suffices, relative to the stronger weights
 456 of faithfulness required to drive the paradigm-uniformity effect.

457 Finally, we compare the joint model to one where we force either the weights of both faith-
 458 fulness constraints to be the same ($\Delta\log\text{-likelihood} = 28.56$, $p < .001$ with one degree of freedom,
 459 $\Delta\text{AICc} = 56.54$, favoring the model with independently-weighted faithfulness constraints), or one
 460 where faithfulness is not scaled by frequency ($\Delta\log\text{-likelihood} = 761.88$, $p < .001$ with one degree

461 of freedom, $\Delta\text{AICc} = 1,466.66$, favoring the model where faithfulness is scaled by frequency). In
462 both cases, the more complex model presented here fits the data significantly better than the
463 alternatives, supporting two basic tenets of the Voting Bases model: multiple faithfulness con-
464 straints active in the grammar, each scaled by the resting activation of the lexical items they refer
465 to (with frequency being the current proxy for resting activation). The R^2 values for the combined
466 model components are within 0.01 of their values for the separately-fit models, demonstrating no
467 compromise in absolute model fit.

468 **6 Discussion**

469 This paper has proposed a model of variable voiced velar nasalization in Japanese, drawing on ex-
470 perimental data published in Breiss et al. (*to appear*). The model integrates grammatical and func-
471 tional determinants of variation, drawing on the Voting Bases framework of lexicon-grammar
472 interaction, which was originally developed to model an entirely separate phonological phe-
473 nomenon, Lexical Conservatism in English and Spanish (Breiss, 2024). Here, we address several
474 major issues that the model raises, notably about whether the proposed system can be learned
475 from the actual Japanese lexicon (section 6.1), about the competence-performance distinction
476 (section 6.2), and about how the Voting Bases model’s mechanism for integrating usage frequency
477 and formal grammar compares to other propositions in the literature (section 6.3). Finally, we
478 close the paper with a more general discussion about how we might understand the broader em-
479 pirical landscape of frequency effects in phonological patterning in light of the proposal in this
480 paper.

481 **6.1 Whence the weights? Evidence in the lexicon**

482 Having observed that there is robust frequency-conditioning of nasalization in both existing and
483 novel compounds, we can ask what the source of this frequency-conditioning might be. By hy-
484 pothesis, the relationship between frequency and resting activation is one that is automatic and
485 not overtly learned. However, we find that the model performs significantly better when al-
486 lowed to set the weights of faithfulness constraints referencing different allomorphs to different
487 weights. This result suggests that, setting aside the relationship between frequency and activa-
488 tion, the speakers must be able to attribute different amounts of influence to different faithfulness
489 constraint violations depending on which base the violation is assessed against. Put another way,
490 the learner needs to be able to figure out how analogically-driven her lexicon is. Here, we present
491 a preliminary investigation of what kind of evidence might exist in the Japanese lexicon that could
492 allow speakers to assign different weights to ID-[nasal]-COMPOUND and ID-[nasal]-N2.

493 We fit a grammar with the constraints in section 4.3 and frequency-driven scaling of faithful-
494 ness violations to the set of compounds in the corpus analyzed by Breiss et al. (2021b) that had
495 a free N2. We found that the optimal weights of the grammar were zero for both *INTERNAL-[g]
496 and ID-[nasal]-COMPOUND, and 1.08 for ID-[nasal]-N2. We had anticipated there being little to
497 no weight assigned to the markedness constraint in this dataset for the same reasons discussed
498 above in section 5.3, but we also found that instead of a tension between faithfulness to the com-
499 pound itself and faithfulness to the N2, the grammar instead left it to the paradigm uniformity
500 effect alone to perturb the otherwise at-chance distribution of variation (at chance because the
501 weight of ID-[nasal]-COMPOUND was at zero, indicating, all else equal, that the alternating and
502 non-alternating candidates were equiprobable). This is qualitatively the same finding as for the
503 novel compounds.

504 We compared the model fit to the corpus data to one where the grammar was forced to assign
505 the same weight to ID-[nasal]-COMPOUND and ID-[nasal]-N2, and found that it was significantly
506 out-performed by the model that allowed the grammar to allot differing weights to different
507 faithfulness constraints to different bases ($\Delta\log\text{-likelihood} = 45.3$, $p < .001$ with one degree of
508 freedom). We take this as tentative evidence that there is an empirical basis in the lexicon for
509 assigning different degrees of faithfulness to different bases.

510 **6.2 Competence, performance, and formal modeling**

511 This paper has proposed a model of Japanese nasalization that integrates token frequency into the
512 workings of the phonological grammar. Since the prospect of integrating a putatively performance-
513 related factor like token frequency into a formal phonological model is not an uncontroversial
514 one, below we directly address some possible criticisms of this approach. We certainly do not
515 think that these are the last words on the topic, but we do feel that by explicitly discussing what
516 we are doing and our motivations for doing it, we take a first step towards a clearer understanding
517 of the stakes and consequences of the choices made in modeling information about usage jointly
518 with the phonological grammar.

519 One initial objection to formally modeling the frequency-conditioned variation in nasaliza-
520 tion might be that there is nothing competence-related to model here at all—the variation is
521 solely driven by “performance” factors (Chomsky, 1965). We respond that this cannot be true
522 of Japanese nasalization: the fact that only compounds whose N2 is morphologically free ex-
523 hibit frequency-sensitive variation, despite the existence of bound morphemes with [g]- and [ŋ]-
524 initial forms like [ga]/[ŋa] “fang”, as shown by the examples in (5), requires an explanation that
525 makes reference to grammatical structures. Further afield, cases like Lexical Conservatism much
526 more strongly blur the line between the contents of the lexicon and the phonological grammar
527 and are well-modeled by a framework like Voting Bases. The fact that this paper demonstrates

528 both paradigm uniformity and Lexical Conservatism emerge as special cases of the same theory
529 speaks to the theoretical insight that can be gained by jointly modeling “performance-related”
530 and “competence-related” influences on the phonological grammar.

531 Another objection to the grammatical model that may be put forward in the current paper
532 is that by incorporating both resting activation (a psycholinguistic quantity) and phonological
533 markedness (a grammatical quantity) into the same formal model, we compromise the distinc-
534 tion between competence and performance to the extent that it is not clear what we can actually
535 say the model is a model *of*. If true, this would indeed be a flaw of the approach; however, a virtue
536 of the Voting Bases model is that lexical influence on the grammar is clearly delimited: the model
537 only allows the lexicon to scale the weights of faithfulness constraints to corresponding lexical
538 representations. Manipulating the resting activation of a given UR has identifiable, localized in-
539 fluences on the computations of the phonological grammar, and instantiates a linking hypothesis
540 consistent with a consensus view of the basic structure of the lexicon. This mechanism can be
541 seen as one way of implementing the idea of “grammar dominance” put forth, for example, by
542 Coetzee (2016) and Coetzee and Kawahara (2013). The “core” phonological grammar—weighted
543 constraints which can assess violations of novel candidates—can be recovered by simply ignor-
544 ing the influence of the lexicon on constraint violations, and can be studied in novel contexts like
545 *wug*-tests, where there is no relevant lexical representation to bear on the grammar.

546 A final objection that we consider is that the very act of jointly modeling usage frequency
547 and the phonological grammar risks leading the analyst to think of fundamentally performance-
548 related factors as in fact competence-related, thus undercutting the goal of researchers whose
549 focus is only understanding linguistic competence. We contend that this is simply false, and in
550 fact, the reverse is true: for a researcher who *only* cares about linguistic competence, modeling
551 usage factors jointly with theories of competence is vital. When confronting data derived from
552 language use (that is, modeling corpus data as in Breiss et al. (2021a), or experimental data where
553 stimuli are existing morphemes of the language as in Breiss et al. (*to appear*), a joint model will
554 better expose the true influence of competence-related factors on the data under study, with the
555 performance-related parts of the model accounting for the otherwise-distorting influence of these
556 factors. Simply ignoring performance-related factors in a formal model makes the strong claim
557 that they have no effect, an assumption which is untenable in the cases examined here, and, we
558 suggest, is also false in many (if not all) types of linguistic data that speakers might have prior
559 usage-based experience with (Arnon and Snider, 2010; Smith and Moore-Cantwell, 2017; Zymet,
560 2018; Morgan and Levy, 2016, 2023). Rather, an integrated approach that jointly models grammar
561 and usage is essential to disentangle competence from performance factors, if this is the goal of
562 the analysis.

6.3 Comparison with other models

The Voting Bases model is one of several approaches in the literature that propose to model the interaction of usage frequency and phonological grammar. In particular, it is similar to the methods proposed in Coetzee and Pater (2008) and Coetzee and Kawahara (2013) which directly scale the weight of faithfulness constraints by the frequency of the form they make reference to, and that of Baird (2021) where a simulated perception-production loop comes to the same result via online learning. This family of approaches involves lowering the weight of faithfulness constraints to high-frequency forms relative to lower-frequency forms which enables them to model data like coronal stop deletion in English (Coetzee and Kawahara, 2013), where higher-frequency monomorphemes (like *just*) tend to get produced more often with a deleted coronal stop than phonologically-similar words (like *jest*). Common to these models is that they assume that the underlying form is /t/-ful, and thus the task of their model must relate higher frequency to therefore have lower constraint weights for it.

A weakness of these models is that, with the possible exception of Baird (2021), the directionality between frequency and constraint weight is arbitrary—the primary goal set in these studies was to fit the data, which is better than the alternative which does not model the effects of lexical frequencies at all, but they suffered somewhat for the lack of clear functional grounding the relation.

By contrast, the frequency-faithfulness relation that Voting Bases model adopts runs in the opposite direction—more frequent forms exact a greater penalty for unfaithful realizations relative to less frequent forms; constraint violations are less severe for low-frequency vs. high-frequency forms. This allows the model to fit a similar range of data, but with a linking hypothesis that is explicitly rooted in resting activation, a construct that is externally justified by a large body of work in psycholinguistics, as reviewed in Breiss (2021, 2024). Lexical items with higher resting activation are more insistent on faithfulness to themselves, corresponding to their increased salience in the language processing system. The main contribution of the Voting Bases model in modeling this phenomenon is that the influence of the lexicon on the grammar should be, in principle, derivable without reference to any facts about the experiment in question; given some independently-established computationally-implemented model of lexical dynamics that represents a scalar quantity of resting activation (or similar construct), the strong prediction of the Voting Bases model is that that quantity should be able to be a fully adequate scaling factor for faithfulness constraint violations. The specific mechanism that is used in this paper—scaling the weights by the sigmoidal transformation of the resting activation—is used since it represents, to us, a reasonable first stab, but the linking function may need to be revised in light of future findings.

In summary, we suggest that the Voting Bases model, because of its functional grounding

599 of frequency effects in externally-motivated psycholinguistic phenomena, is on firmer footing
600 than theories that have alternative linking functions between frequency and grammar, which are
601 arguably arbitrary.

602 **6.4 Towards a unified picture of token frequency in phonology**

603 In this section, we broaden our view of token frequency effects in phonology, and discuss how
604 considering the varying functional roles of frequency can reconcile some seemingly-contradictory
605 bodies of evidence.

606 First, there is evidence that higher token-frequency leads to more markedness-reducing al-
607 ternations. Coetzee and Kawahara (2013) found that higher-frequency lexical items were more
608 likely to undergo phonological processes of simplification and (markedness-)reduction: high-
609 frequency English words like *jus(t)* underwent an optional process of coronal stop deletion at a
610 higher rate than low-frequency words like *jes(t)*, and high-frequency Japanese words like [baggu]
611 “bag” underwent geminate devoicing more often than low-frequency words like [budda] “Bud-
612 dha” (Kawahara and Sano, 2013). Zuraw (2007) examines frequency-conditioned application of
613 markedness-reducing phonological processes in a corpus of written Tagalog, and likewise finds
614 higher rates of repair within higher-frequency units (words, clitic groups, etc), subject to the
615 markedness principles of the language.

616 On the other hand, there is also evidence to show that higher-frequency forms are more
617 likely to be exceptional, and thus marked with regard to the overall properties of the grammar.
618 Smith and Moore-Cantwell (2017) found that higher-frequency comparative constructions are
619 more likely to flout grammar-wide trends driven by markedness. In a similar vein, Anttila (2006)
620 and Mayer (2021) found that higher-frequency morphologically-complex forms were more likely
621 to behave opaquely with respect to grammar-wide phonological processes.

622 We can compare these effects to the ones observed in Breiss et al. (2021b) and Breiss et al.
623 (*to appear*): higher-frequency N2s act as stronger attractors, yielding *more* faithfulness to their
624 preserved surface [g] resulting in lower rates of nasalization, whereas higher compound fre-
625 quency as a whole yielded higher rates of nasalization. Thus it seems that for compounds, higher
626 frequency is correlated with more phonological-process application and markedness-reduction;
627 this is broadly in line with the findings of Coetzee and Kawahara (2013) where higher-frequency
628 words undergo more phonological alternations. However, we found that at the same time, in
629 compounds with free N2s, higher free N2 frequency is related to *less* process application, with
630 higher-frequency supporting the retention of a marked structure (word-medial [g]).

631 We suggest we can resolve this tension by distinguishing between the processes that token
632 frequency can impact: one is whether to set up an independent lexical representation for a surface
633 allomorph, and the other is influencing the strength of that representation in the lexicon of the

634 speaker.

635 If a form is exceptional and high-frequency, it may be more economical for a speaker to pay a
636 one-time “cost” of encoding the exception as a listed form that is not derived by the grammar, thus
637 relieving the phonology of the difficulty of having to generate the exceptional or idiosyncratic
638 form on each of the many frequent occasions of use (cf. Adaptor Grammars (Johnson et al., 2007,
639 *et seq.*) or Fragment Grammars (O’Donnell, 2015) which offer computationally-explicit imple-
640 mentations of this general idea). For lower-frequency exceptional forms, the likelihood of listing
641 is less since the price trades off less favorably with the amount of times it is used; thus lower-
642 frequency forms are more susceptible to change and regularization to the dominant grammatical
643 trends over time compared to higher-frequency forms. Another aspect of this trade-off is the
644 emergence of Lexicon Optimization (Prince and Smolensky, 1993; Sanders, 2003, 2006); even if a
645 form is not particularly exceptional, if a UR almost always surfaces with a phonological process
646 applied to it, with sufficient frequency it becomes less costly to just store the form with phonolog-
647 ical process applied—that is, create a separate allomorph that is specific to the environment that
648 would trigger the phonological rule. This, similarly, relieves the grammar of the job of having to
649 repair the form every time. Thus, we find Lexicon Optimization targeting forms like *jus(t)* over
650 forms like *jest*, making these forms restructured to automatically have the phonological alterna-
651 tion applied, thus giving the appearance of having undergone a markedness-improving repair in
652 the grammar, but actually the frequency of the form has resulted in restructuring to the lexicon
653 (see Breiss and Wilson (2020) for an initial attempt at a computational model of the phonological
654 grammar and lexicon that exhibits this property).

655 As reviewed above, lexical frequency also influences the resting activation of a lexical item
656 once it is listed in the lexicon. In the Voting Bases model, higher resting activation leads to
657 the listed form exerting a stronger pull on the surface realization of a related form; where this
658 pressure goes against the broader principle of markedness in the grammar, as in cases of paradigm
659 uniformity, we find that marked structures with high-frequency output-bases are preserved; in
660 cases where the listed form coincides with the output of the markedness-reducing process, as
661 in many cases of Lexical Conservatism (Steriade, 1997; Steriade and Stanton, 2020; Breiss, 2021),
662 then the higher-frequency form promotes an unmarked surface form.

663 Recent work by Jarosz et al. (2024) has laid out a class of models which exhibit characteristics
664 that align favorably with the dynamics of frequency laid out here, suggesting that an integrated,
665 implemented model that can jointly account for the variety of frequency effects reviewed in this
666 section is perhaps quite close at hand. Future work may profitably explore how well these mod-
667 els can provide converging evidence from computational learning simulations to support the
668 psycholinguistic, experimental, and diachronic evidence for the contents of the lexicon that the
669 Voting Bases theory relies on. In sum, the broader landscape of token frequency in phonology

670 is compatible with the functional grounding given to frequency under the Voting Base model,
671 though much empirical and formal work remains to be done to further support the predictions of
672 the framework more broadly as a candidate for a general theory of the influence of the dynamic
673 lexicon on the probabilistic grammar.

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